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FINC 520, Spring 2007  
Homework 3, due Wednesday, April 25.

1. Consider the four variable VAR you estimated for the previous homework. Define the ‘business cycle frequencies’ as the component of data corresponding to frequencies of fluctuation between 1 and 8 years (i.e., 4 and 32 quarters). Let ‘high frequencies’ denote the component of data corresponding to fluctuations between 2 quarters and 1 year. Let ‘low frequencies’ denote the component corresponding to fluctuations longer than 8 years.
  - (a) For the short rate, the predicted (i.e., using the expectations hypothesis) long rate and the actual short rate, compute and display the share of variance in the low frequencies, the business cycle frequencies and the high frequencies. How well does the expectations hypothesis do in predicting the variance decomposition, in frequency domain, of the long rate? What is the variance of the long rate predicted by the expectations hypothesis, divided by the variance of the short rate?
  - (b) Instead of using the  $\phi_1, \phi_2, \phi_3, \phi_4$  computed from the data, use  $\phi_1 = \phi_2 = \phi_3 = \phi_4 = 0$ . Thus, all the data are now a white noise (about a constant). What is the effect of this change on the distribution of variance of the short and predicted long rates across different frequency bands? What is the effect of this change on the relative variance of the predicted long rate to the variance of the short rate? Explain, using economic intuition.
  - (c) Now set  $\phi_2 = \phi_3 = \phi_4 = 0, \phi_1 = 0.999 \times I$ . Repeat the calculations in (b) above. Explain your results, using economic intuition.
2. Suppose that two time series processes are related as follows:

$$y_t = \sum_{k=-\infty}^{\infty} h_k x_{t-k} + \varepsilon_t,$$

where all variables have mean zero and  $\varepsilon_t$  is uncorrelated with  $x_{t-j}$  for all  $j$ . Evidently, the above relationship is the projection of  $y_t$  onto

$\{\dots, x_{-1}, x_0, x_1, \dots\}$ . Let

$$\Gamma_{yx,k} \equiv Ey_t x_{t-k}, \quad \Gamma_{x,k} \equiv Ex_t x_{t-k},$$

and show that

$$g_{yx}(e^{-i\omega}) = h(e^{-i\omega}) g_x(e^{-i\omega}),$$

where

$$g_{yx}(e^{-i\omega}) = \sum_{k=-\infty}^{\infty} \Gamma_{yx,k} e^{-i\omega k}, \quad g_x(e^{-i\omega}) = \sum_{k=-\infty}^{\infty} \Gamma_{x,k} e^{-i\omega k}, \quad h(e^{-i\omega}) = \sum_{k=-\infty}^{\infty} h_k e^{-i\omega k}.$$

Conclude that the Fourier transform of the projection coefficients is given by:

$$h(e^{-i\omega}) = \frac{g_{yx}(e^{-i\omega})}{g_x(e^{-i\omega})},$$

so that the individual coefficients may be recovered from the relation:

$$h_k = \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{g_{yx}(e^{-i\omega})}{g_x(e^{-i\omega})} e^{i\omega k} d\omega,$$

which may be approximated by a Riemann sum.

3. (Optimal seasonal adjustment). Suppose that a time series,  $\{X_t\}$ , has the following representation:

$$X_t = x_t + u_t,$$

where  $x_t$  and  $u_t$  are purely indeterministic and ergodic, covariance stationary processes. Suppose that  $u_t$  is the source of seasonality in  $X_t$ . That is, the spectrum of  $u_t$ ,  $S_u(e^{-i\omega})$ , has much power concentrated in the seasonal frequencies (i.e., those near  $\omega = 2\pi/4$  in quarterly data). The spectrum of  $x_t$ ,  $S_x(e^{-i\omega})$ , is smooth and does not display a peak in the seasonal frequencies. The econometrician seeks to estimate the ‘seasonally adjusted data’,  $x_t$ , by projecting  $x_t$  onto a complete realization of  $X_t$  (i.e.,  $\{\dots, X_{-1}, X_0, X_1, \dots\}$ ):

$$x_t = \sum_{k=-\infty}^{\infty} h_k X_{t-k} + v_t,$$

where  $v_t$  is uncorrelated with  $X_{t-k}$  for all  $k$ . Let  $\hat{x}_t$  denote the ‘seasonally adjusted’ data:

$$\hat{x}_t = \sum_{k=-\infty}^{\infty} h_k X_{t-k}.$$

- (a) Derive the formula for  $h(e^{-i\omega})$  in terms of the known objects,  $S_u(e^{-i\omega})$  and  $S_x(e^{-i\omega})$ .
  - (b) Show that  $S_{\hat{x}}(e^{-i\omega}) < S_x(e^{-i\omega})$  for all  $\omega$ .
  - (c) Show that if  $S_x(e^{-i\omega})$  is smooth across all frequencies, while  $S_u(e^{-i\omega})$  has sharp peaks at the seasonal frequencies, then  $S_{\hat{x}}(e^{-i\omega})$  will have substantial *dips* at the seasonal frequencies. (It is ironic that optimal seasonal adjustment produces a series,  $\hat{x}_t$ , that itself displays seasonality.)
4. (Hodrick-Prescott filter). Suppose we a partial realization,  $y_1, \dots, y_T$ , from a covariance stationary, indeterministic and ergodic time series,  $\{y_t\}$ . The HP filter solves the problem:

$$\min_{\{y_t^T\}} \sum_{t=1}^{T-1} \left\{ (y_t - y_t^T)^2 + \lambda \left[ (y_{t+1}^T - y_t^T) - (y_t^T - y_{t-1}^T) \right]^2 \right\},$$

and the ‘HP-filtered’ data are

$$y_t^c \equiv y_t - y_t^T.$$

Suppose  $T$  is large and the start of the data set is arbitrarily far in the past.

- (a) Construct the filter,  $g(L)$ , which expresses

$$y_t^c = g(L) y_t,$$

where  $g(L)$  is called ‘the HP filter’. (Hint: first compute the first order necessary condition for optimality satisfied by  $y_t^T$ . In lag operator form, this has the representation,

$$y_t = B(L) y_t^T,$$

where  $B(L)$  is symmetric in positive and negative powers of  $L$ , i.e.,  $B(L) = B(L^{-1})$ . Note that since  $y_t^T = y_t - y_t^c$ , this implies:

$$y_t = B(L)y_t - B(L)y_t^c,$$

or,

$$y_t^c = g(L)y_t,$$

where the *HP* filter,  $g(L)$  is

$$g(L) = \frac{B(L) - 1}{B(L)}.$$

End of hint!)

- (b) To see what the *HP* filter does to a time series, graph  $g(e^{-i\omega})$  for  $\omega \in (0, \pi)$ . Note that it looks like a ‘high pass filter’. That is, it looks like a band pass filter which lets higher frequencies of oscillation through and zeros out the lower frequencies. What is the (approximate) cutoff between frequencies allowed through and frequencies set to zero? Here, imagine you are working with quarterly data and  $\lambda = 1600$ .

5. Suppose that the data are generated by a true (scalar) autoregressive representation of the following form:

$$y_t = \phi(L)y_{t-1} + \varepsilon_t,$$

where  $\phi(L)$  is a polynomial in non-negative powers of  $L$  and the polynomial coefficients are square-summable. Also,  $\varepsilon_t$  is a white noise, uncorrelated with  $y_{t-s}$ ,  $s > 0$ . Suppose the econometrician estimates  $\phi(L)$  by running a regression of  $y_t$  on  $p$  lags of itself. The econometrician is assumed to have an entire (i.e., doubly infinite) realization of data. The econometrician may commit some form of specification error, for example by choosing a value of  $p$  smaller than the true value (the true lag length may actually be infinite). By ‘running a regression’, the econometrician is assumed to choose coefficients,  $\hat{\phi}_1, \dots, \hat{\phi}_p$ , for the AR polynomial,  $\hat{\phi}(L)$ , so that

$$y_t - \hat{\phi}(L)y_{t-1}$$

has the smallest possible variance in the (infinite!) sample.

- (a) Argue carefully (be clear when you use ergodicity and covariance stationarity) that the econometrician's choice of  $\hat{\phi}(L)$  solves

$$\min_{\hat{\phi}_1, \dots, \hat{\phi}_p} \frac{1}{2\pi} \int_{-\pi}^{\pi} [\phi(e^{-i\omega}) - \hat{\phi}(e^{-i\omega})] g_y(e^{-i\omega}) [\phi(e^{i\omega}) - \hat{\phi}(e^{i\omega})]' d\omega,$$

where  $g_y(e^{-i\omega})$  is the covariance generating function of  $\{y_t\}$ . Note that if the econometrician commits no specification error, then

$$\phi(e^{-i\omega}) = \hat{\phi}(e^{-i\omega}), \text{ for all } \omega \in (-\pi, \pi).$$

- (b) Suppose the econometrician does commit specification error, so that the previous equality is not possible over all frequencies,  $\omega \in (-\pi, \pi)$ . Suppose the econometrician is particularly interested in the sum of the AR coefficients,  $\phi(1)$ . Explain why the econometrician's estimator of this object,  $\hat{\phi}(1)$ , is likely to be a good one if there is an important low-frequency component in the data,  $\{y_t\}$ . Alternatively, if the data are primarily driven by high frequency components, then  $\phi(1)$  is likely to be badly estimated.