

Assessing Structural VARs

by

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- Another tradition:
 - Build macro models with large numbers of shocks - complete characterization of DGP.
 - Avoids KP conundrum.

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 - This is at Heart of Difficulty in With Long-Run Identification
 - See also Faust and Leeper and Pagan.

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 - This is at Heart of Difficulty in With Long-Run Identification
 - See also Faust and Leeper and Pagan.
- More recently EGG and CKM examine reliability of VAR-based inference using long run identifying restrictions.
 - CKM are exceedingly critical.

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 - Look at Long Run Restrictions and Short Run Restrictions

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- We address these questions using data generated from dynamic GE models.
 - Look at Long Run Restrictions and Short Run Restrictions

- Our conclusion:
 - Structural VARs provide valuable information for building empirically plausible models of aggregate fluctuations.

Findings for Short Run Restrictions

- DGP: variants of a standard real business cycle model augmented by timing restrictions..
 - Focus on response of hours to technology shock.
- Conclusion:
 - VAR's perform *remarkably* well
 - Virtually no bias - either in point estimates or estimates of sampling uncertainty.
- Very comforting for vast literature that uses short run restrictions to identify consequences of shocks to economy.

Findings for long run restrictions

- When technology shocks account for a substantial fraction of business cycle fluctuations in output, VAR based analysis is reliable.
 - Some evidence of bias when tech shocks play much smaller role relative to estimates in standard RBC literature.
- First way to eliminate bias:
 - When number of variables in VAR exceeds number of important driving shocks, bias in impulse response estimators is substantially reduced.
 - * Widespread consensus: only a handful (e.g., 3-4) of important shocks drive aggregate fluctuations
- Second way to eliminate bias:
 - Integrate Newey-West non-parametric estimator of zero-frequency spectral density

Outline

- RBC Models
 - Various Parameterizations Considered
 - Standard Version (Long Run Restrictions)
 - Recursive Version (Short Run Restrictions)
- Structural VAR and the Identification Problem
- Short Run Restrictions Approach to Identification
- Long Run Restrictions Approach to Identification
- Reconciling with CKM
- Concluding Comments

DGP: A Generic RBC Model

- Preferences:

$$E \left\{ \sum_{t=0}^{\infty} (\beta (1 + \gamma))^t \left[\log c_t + \psi \frac{(\bar{l} - l_t)^{1-\sigma}}{1-\sigma} \right] \middle| \Omega_0 \right\}.$$

- Constraints:

$$c_t + (1 + \tau_{x,t}) [(1 + \gamma) k_{t+1} - (1 - \delta) k_t] \leq (1 - \tau_{lt}) w_t l_t + r_t k_t.$$

$$c_t + (1 + \gamma) k_{t+1} - (1 - \delta) k_t \leq k_t^\alpha (Z_t l_t)^{1-\alpha}.$$

- Shocks:

$$\log(Z_t) = \mu_Z + \log(Z_{t-1}) + \sigma_z \varepsilon_t^z,$$

$$\tau_{lt+1} = (1 - \rho_l) \bar{\tau}_l + \rho_l \tau_{lt} + \sigma_l \varepsilon_{t+1}^d,$$

$$\tau_{xt+1} = (1 - \rho_x) \bar{\tau}_x + \rho_x \tau_{xt} + \sigma_x \varepsilon_{t+1}^x.$$

Parameterizing the Model

- As in CKM we assume

$$\begin{aligned}\beta &= 0.9722^{1/4}, \theta = 0.35, \delta = 1 - (1 - .0464)^{1/4}, \\ \psi &= 2.24, \gamma = 1.015^{1/4} - 1, \bar{l} = 1300, \\ \bar{\tau}_x &= 0.3, \bar{\tau}_l = 0.27388, \mu_z = 1.016^{1/4} - 1, \sigma = 1.\end{aligned}$$

- Different versions of the RBC model, distinguished by the nature of exogenous shocks.

Parameterizing the Model ...

KP Specification

- Technology shock process (Prescott (1986))

$$\log z_t = \mu_Z + 0.011738 \times \varepsilon_t^z.$$

- EGG (2005) update Prescott's analysis, estimate σ_z to be 0.0148.
 - To be conservative, we use Prescott's estimate.
- Law of motion for $\tau_{l,t}$ as follows.
 - Household / firm FONC's imply:

$$\tau_{l,t} = 1 - \frac{c_t}{y_t} \frac{l_t}{\bar{l}} \frac{\psi}{1 - \theta}.$$

$$\tau_{l,t} = (1 - 0.9934) \times 0.2660 + 0.9934 \times \tau_{l,t-1} + .0062 \times \varepsilon_t^l.$$

- Percent of variance in HP-filtered, log output due to technology shocks is 73%.
 - Consistent with key claim of KP.

Parameterizing the Model ...

CKM Benchmark Specification

$$\begin{aligned}\log z_t &= \mu_Z + \log z_t = \mu_Z + 0.00581 \times \varepsilon_t^z \\ \tau_{lt} &= (1 - \rho_l) \bar{\tau}_l + \rho_l \tau_{l,t-1} + 0.00764 \times \varepsilon_t^l, \quad \rho_l = 0.93782.\end{aligned}$$

- Percent of variance in HP-filtered, log output due to technology shocks is only 23%.
- Irony:
 - CKM use this specification to criticize Gali (1999).
 - Embodies Gali's main hypothesis that technology shocks play only a very small role in business cycle fluctuations.

Parameterizing the Model ...

Other Specifications

- Vary σ and σ_l
 - Important quantitative effect on contribution of technology shocks to volatility of output.
- *Three Shocks, Two Important Specification*
 - Additional (Unimportant) Shock, Capital Tax Rate

$$\tau_{xt} = \bar{\tau}_x + 0.0001 \times \varepsilon_t^x$$

- Three Variables in VAR Analysis:

$$a_t \equiv \log \left(\frac{y_t}{l_t} \right), \log l_t, \log \left(\frac{c_t}{y_t} \right)$$

Parameterizing the Model ...

- *Three Shocks, Three Important Specification*
 - As in CKM:

$$\tau_{xt} = (1 - 0.9)\bar{\tau}_x + 0.9 \times \tau_{x,t-1} + 0.01\varepsilon_t^x.$$

- Three Variables in VAR Analysis

- *Four Shocks, Three Important Specification*
 - Capital Tax as in CKM
 - Four Variables in VAR Analysis

$$a_t \equiv \log\left(\frac{y_t}{l_t}\right), \log l_t, \log\left(\frac{c_t}{y_t}\right), \tau_{xt} + w_t$$

$$w_t \sim N(0, 0.0001)$$

Two Versions of Model

- Differentiated by timing assumptions.
- *Standard version*
 - All time t decisions taken after realization of the time t shocks.
- *Recursive version*
 - First, τ_{lt} is observed. Then, labor decision made.
 - Second, other shocks are realized.
 - Then, agents make their investment and consumption decisions.
 - Finally, labor, investment, consumption, and output occur

Estimating the Effects of a Positive Technology Shock in VAR

- VAR:

$$X_t = B_1 X_{t-1} + B_2 X_{t-2} + \dots + B_p X_{t-p} + u_t,$$

$$Eu_t u_t' = V, u_t = C e_t, E e_t e_t' = I, C C' = V$$

$$X_t = \begin{pmatrix} \Delta \log a_t \\ \log l_t \\ x_t \end{pmatrix}, C = [C_1 : C_2 : C_3], \varepsilon_t = \begin{pmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \\ \varepsilon_{3t} \end{pmatrix}, a_t = \frac{Y_t}{l_t}$$

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- Impulse Response to Positive Technology Shock (Say, $\varepsilon_{1t} = \varepsilon_t^z$):

$$\begin{aligned} X_t - E_{t-1} X_t &= C_1 \varepsilon_{1t}, E_t X_{t+1} - E_{t-1} X_{t+1} = B_1 C_1 \varepsilon_{1t} \\ E_t X_{t+2} - E_{t-1} X_{t+2} &= B_1^2 C_1 \varepsilon_{1t} + B_2 C_1 \varepsilon_{1t} \end{aligned}$$

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- Need: B_1, \dots, B_p, C_1 .

Identification Problem

- From Applying OLS To Both Equations in VAR, We ‘Know’:

$$B_1, \dots, B_p, V$$

- Problem, Need first Column of C , C_1
- Restrictions (Bivariate Case): three equations in four unknowns

$$CC' = V$$

- Identification Problem:

Not Enough Restrictions to Pin Down C_1

- Need More Restrictions

The Recursive Version of the Model

- First, τ_{lt} is observed. Then, labor decision made. Consequently,

$$u_{\Omega,t}^l = \gamma \varepsilon_t^l, \quad u_{\Omega,t}^l \equiv P[l_t | \Omega_{t-1}]$$

- Second, other shocks are realized, so

$$u_{\Omega,t}^a = \alpha_1 \varepsilon_t^z + \alpha_2 \varepsilon_t^l, \quad u_{\Omega,t}^a \equiv P[a_t | \Omega_{t-1}]$$

- Regression:

$$u_{\Omega,t}^a = \beta u_{\Omega,t}^l + \alpha_1 \varepsilon_t^z, \quad \beta = \frac{\text{cov}(u_{\Omega,t}^a, u_{\Omega,t}^l)}{V(u_{\Omega,t}^l)},$$

- Perform Analogous Calculations in VAR

The Recursive Version of the Model ...

$$Y_t = \begin{pmatrix} \log l_t \\ \Delta \log a_t \\ x_t \end{pmatrix},$$

$$u_t = \begin{pmatrix} u_t^l \\ u_t^a \\ u_t^x \end{pmatrix}$$

$$u_t = C\varepsilon_t, E\varepsilon_t\varepsilon_t' = I, CC' = V$$

$$\varepsilon_{2t} \sim \varepsilon_t^z$$

- For Response of Y_t to ε_t^z , need B_1, \dots, B_q and second column of C .
 - Compute $CC' = V$, where C is lower triangular Choleski decomposition of V .
 - Take second column of C .
- Potential Source of Specification Error: Differences Between One-Step-Ahead Forecast Errors in Model and VAR.

Experiments

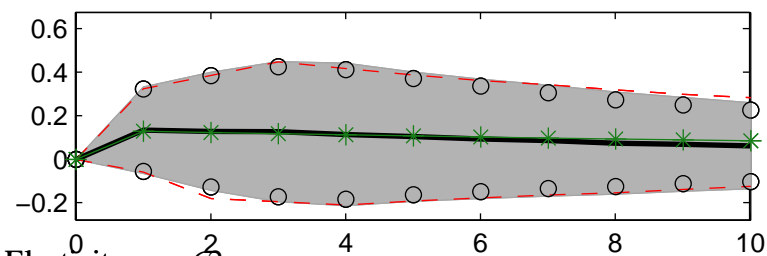
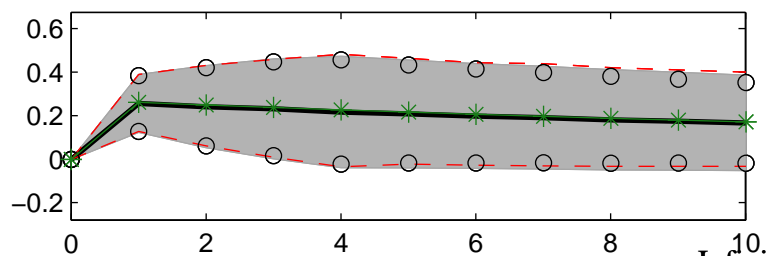
- Simulate 1000 data sets, each of length 180 observations, using GE model as DGP.
 - Shocks ε_t^z , ε_t^l and possibly ε_t^x are drawn from *i.i.d.* standard normal distributions.
- Estimate a four lag VAR.
 - Report Mean Impulse Response Function over 1000 synthetic data sets.
 - Measure of sampling uncertainty associated with the estimated dynamic response functions.
 - * Calculate standard deviation of points in estimated impulse response functions across the 1000 synthetic data sets (Grey Area).
 - * Also calculate middle 95% of the estimated coefficients in dynamic response functions across the 1000 synthetic data sets (Red lines).
 - Report Mean of Econometrician's Confidence Interval

Figure 2: Analysis of Short-Run Identification Assumption

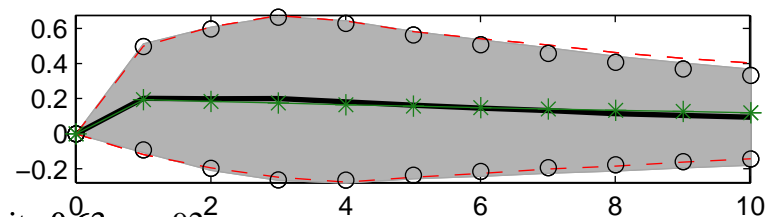
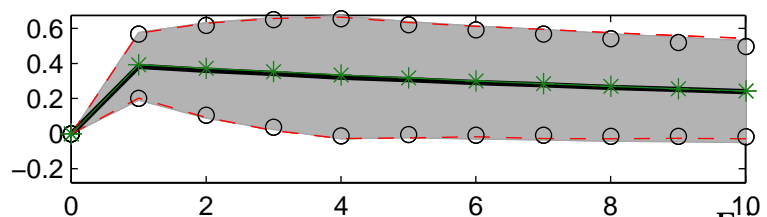
KP Model

CKM Model

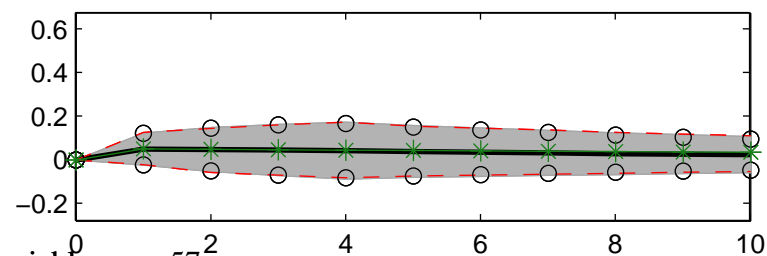
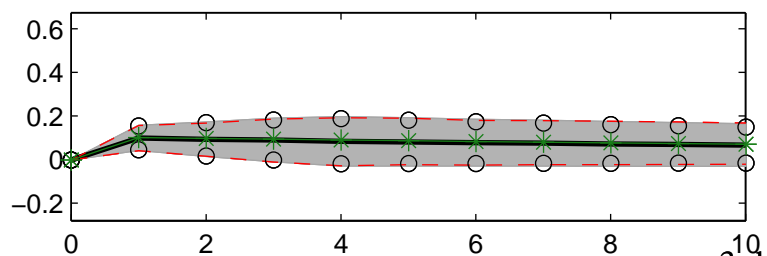
Benchmark, $p=73$



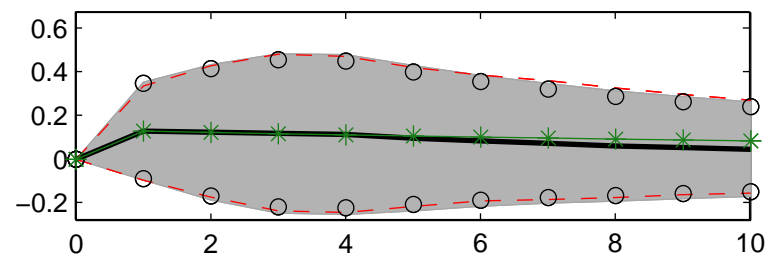
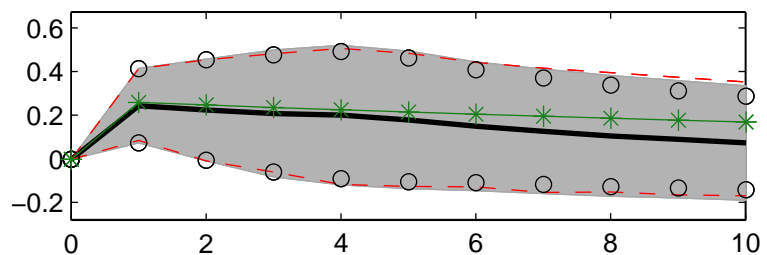
Infinite Frisch Elasticity, $p = 62$



Frisch Elasticity 0.63, $p = 92^2$



3 shocks, 3 variables, $p = 57^2$



Note: hours response, in percent terms, to a 1.2 (KP) or 0.6 (CKM) percent innovation in technology, Z_t .

Solid line – mean response, Gray area – mean response plus/minus two standard errors,
 Starred line – true response, Dashed line – 95.5 percent probability interval of responses,
 Circles – average value of econometrician estimated plus/minus two standard errors.

Summary of Findings with Short Run Restrictions

- No evidence of bias in the estimated impulse response functions.
- An econometrician wouldn't be misled in inference using standard procedures for constructing confidence intervals.
- SVAR's perform remarkably well.
 - Absent specification error, standard structural VAR procedures reliably uncover and identify the dynamic effects of shocks to the economy.
- We did *not* include capital as a variable in the VAR.
 - Claims in CKM to contrary, omitting economically relevant state variable capital does not in and of itself pose a problem for inference using structural VAR's.

Long-Run Restrictions

- Two Key Properties of Model:
 - Exclusion Restriction:

$$\lim_{j \rightarrow \infty} [E_t a_{t+j} - E_{t-1} a_{t+j}] = f(\varepsilon_t^z \text{ only})$$

- Sign Restriction:

f increasing in ε_t^z

- Exploit Analogous Properties in VAR to Identify Technology Shocks and their Effects

Applying Analogous Restrictions to VAR

- Note:

$$\begin{aligned}\tilde{E}_t[a_{t+1}] - \tilde{E}_{t-1}[a_{t+1}] &= \tilde{E}_t[\Delta a_{t+1} + \Delta a_t] - \tilde{E}_{t-1}[\Delta a_{t+1} + \Delta a_t] \\ &= [\tilde{E}_t \Delta a_{t+1} - \tilde{E}_{t-1} \Delta a_{t+1}] + [\Delta a_t - \tilde{E}_{t-1} \Delta a_t]\end{aligned}$$

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- Then ($p = 1$)

$$\tilde{E}_t[a_{t+1}] - \tilde{E}_{t-1}[a_{t+1}] = (1, 0) [B + I] C \varepsilon_t$$

$$\tilde{E}_t[a_{t+2}] - \tilde{E}_{t-1}[a_{t+2}] = (1, 0) [B^2 + B + I] C \varepsilon_t$$

$$\tilde{E}_t[a_{t+j}] - \tilde{E}_{t-1}[a_{t+j}] = (1, 0) [B^j + B^{j-1} + \dots + B^2 + B + I] C \varepsilon_t$$

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$$\tilde{E}_t[a_{t+j}] - \tilde{E}_{t-1}[a_{t+j}] = (1, 0) [B^j + B^{j-1} + \dots + B^2 + B + I] C \varepsilon_t$$

as $j \rightarrow \infty$:

$$\begin{aligned} & \lim_{j \rightarrow \infty} \tilde{E}_t[a_{t+j}] - \tilde{E}_{t-1}[a_{t+j}] \\ &= \lim_{j \rightarrow \infty} (1, 0) [\dots + B^j + B^{j-1} + \dots + B^2 + B + I] C \varepsilon_t \\ &= (1, 0) [I - B]^{-1} C \varepsilon_t \end{aligned}$$

Applying Analogous Restrictions to VAR ...

- As $j \rightarrow \infty$ (for arbitrary p) :

$$\lim_{j \rightarrow \infty} \tilde{E}_t[a_{t+j}] - \tilde{E}_{t-1}[a_{t+j}] = (1, 0, \dots, 0) [I - B(1)]^{-1} C \varepsilon_t$$

$$B(1) = B_1 + B_2 + \dots + B_p$$

- $\tilde{E}_t \sim$ Expectation, Conditional on Information Set in VAR
 - Potential Specification Error
 - * Too Few Variables in VAR
 - * Too Few Lags in VAR

Applying Analogous Restrictions to VAR ...

- The VAR:

$$X_t = B_1 X_{t-1} + B_2 X_{t-2} + \dots + B_p X_{t-p} + u_t$$

- Identification: Solve for C Such that -

$$\text{(exclusion restriction) } [I - B(1)]^{-1} C = \begin{bmatrix} \text{number} & 0, \dots, 0 \\ \text{numbers} & \text{numbers} \end{bmatrix}$$

(sign restriction) (1, 1) element of $[I - B(1)]^{-1} C$ is *positive*

$$CC' = V$$

- There Are Many C That Satisfy These Constraints. All Have the Same C_1 .

Standard Algorithm for Computing C_1

- Step 1: Compute Lower Triangular Choleski Decomposition, D

$$DD' = [I - B(1)]^{-1} V [I - B(1)']^{-1} = S_0 \text{ ('Spectral Density of } X_t \text{ at Frequency Zero')}$$

subject to $D(1, 1) > 0$.

- Step 2: Solve

$$C = [I - B(1)] D.$$

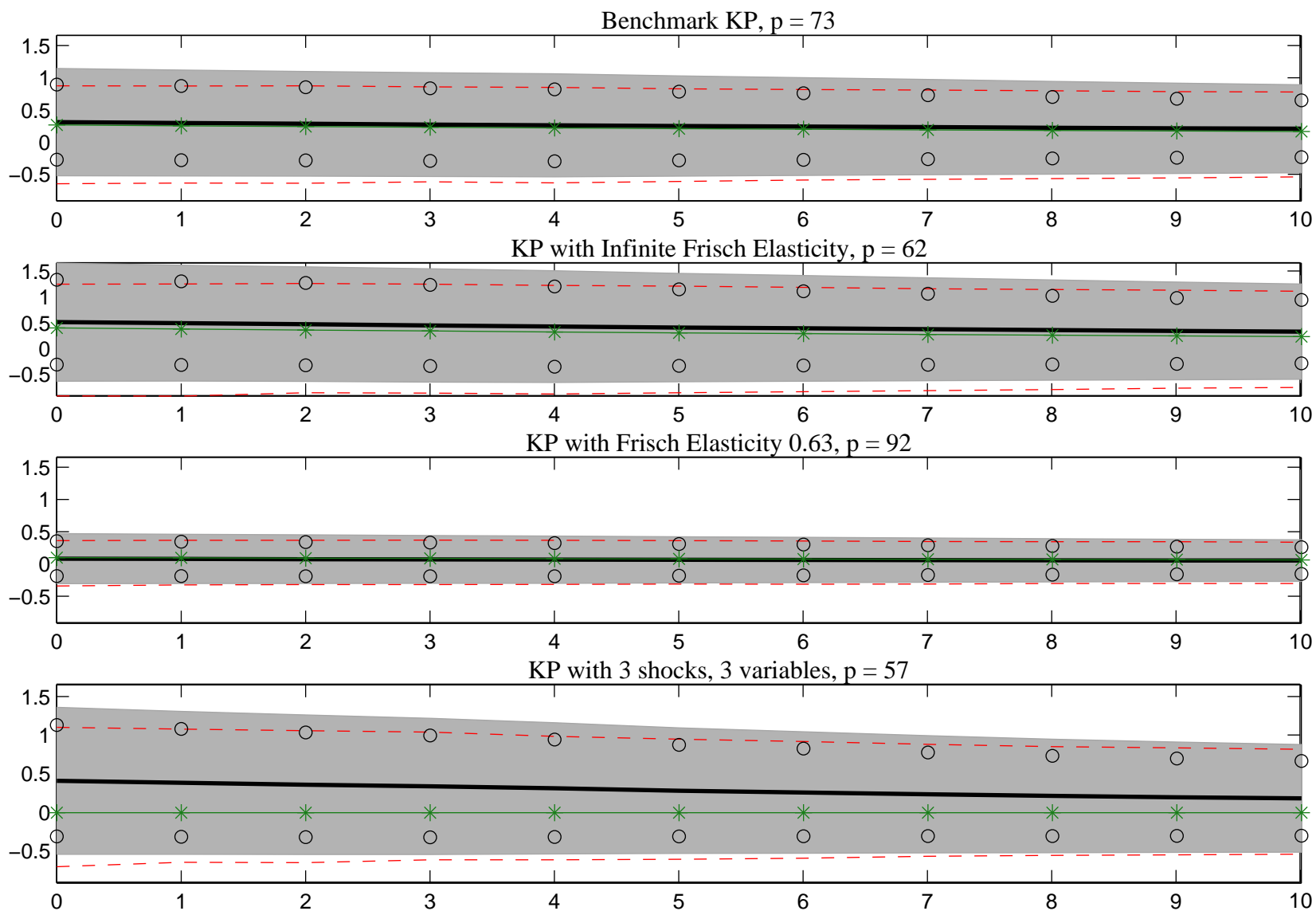
- Remark: this C Satisfies all Restrictions

$$CC' = [I - B(1)] DD' [I - B(1)'] = V$$

$$\text{(exclusion restriction) } [I - B(1)]^{-1} C = \begin{bmatrix} x & 0, \dots, 0 \\ \text{numbers} & \text{numbers} \end{bmatrix}$$

$$\text{(sign restriction) } x > 0$$

**Figure 3a: Analysis of the Long-Run Identification Assumption with Kydland-Prescott Specification
Standard Estimator**



Note: hours response, in percent terms, to a 1.2 percent innovation in technology, Z_t .

Solid line – mean response, Gray area – mean response plus/minus two standard errors,
Starred line – true response, Dashed line – 95.5 percent probability interval of responses,
Circles – average value of econometrician estimated plus/minus two standard errors.

Long Run Restrictions: KP Specification

- Virtually no bias in point estimates.
- Considerable sampling uncertainty, but econometrician wouldn't be misled with respect to inference.
- Hansen Indivisible Labor model, $\sigma = 0.0001$.
 - Bias associated with estimator increases (very) slightly.
 - * Percent of variance in HP-filtered, log output due to technology shocks is 62%.
 - Econometrician wouldn't be misled about sampling uncertainty.
- EGG: $\sigma = 1.24$ (Frisch elasticity = 0.63)
 - Bias almost disappears, and the sampling uncertainty shrinks drastically.
 - Percent of variance in HP-filtered, log output due to technology shocks is 92%.

Long Run Restrictions: KP Specification ...

- Three variable, three shock version of model.
 - Noticeable degree of bias associated with the estimated impulse response function.
 - * But relatively small in relation to the sampling variation.
 - * Econometrician's estimated confidence interval is roughly correct, on average.
 - * Percent of variance in HP-filtered, log output due to technology shocks is 57%

Why Does Bias Appear in Last Case?

- Sims (1972) : can characterize the VAR parameter estimates econometrician would obtain in large sample ($\hat{B}_1, \dots, \hat{B}_q$ and \hat{V})

$$\hat{V} = V + \min_{\hat{B}_1, \dots, \hat{B}_q} \frac{1}{2\pi} \int_{-\pi}^{\pi} \left[B(e^{-i\omega}) - \hat{B}(e^{-i\omega}) \right] S_Y(\omega) \left[B(e^{i\omega}) - \hat{B}(e^{i\omega}) \right]' d\omega$$

– $S_Y(\omega)$ is associated spectral density, at frequency ω .

- Econometrician chooses VAR lag matrices to minimize a quadratic form in difference between estimated and true lag matrices
 - Assigns greatest weight to frequencies where spectral density is greatest.
 - If there's specification error, then $\hat{B} \neq B$ and $V > \hat{V}$.
- Specification error:
 - Model Implies $q = \infty$, But Econometrician uses $q = 4$.
 - Model May Call for More Variables in Analysis.

Why Does Bias Appear in Last Case? ...

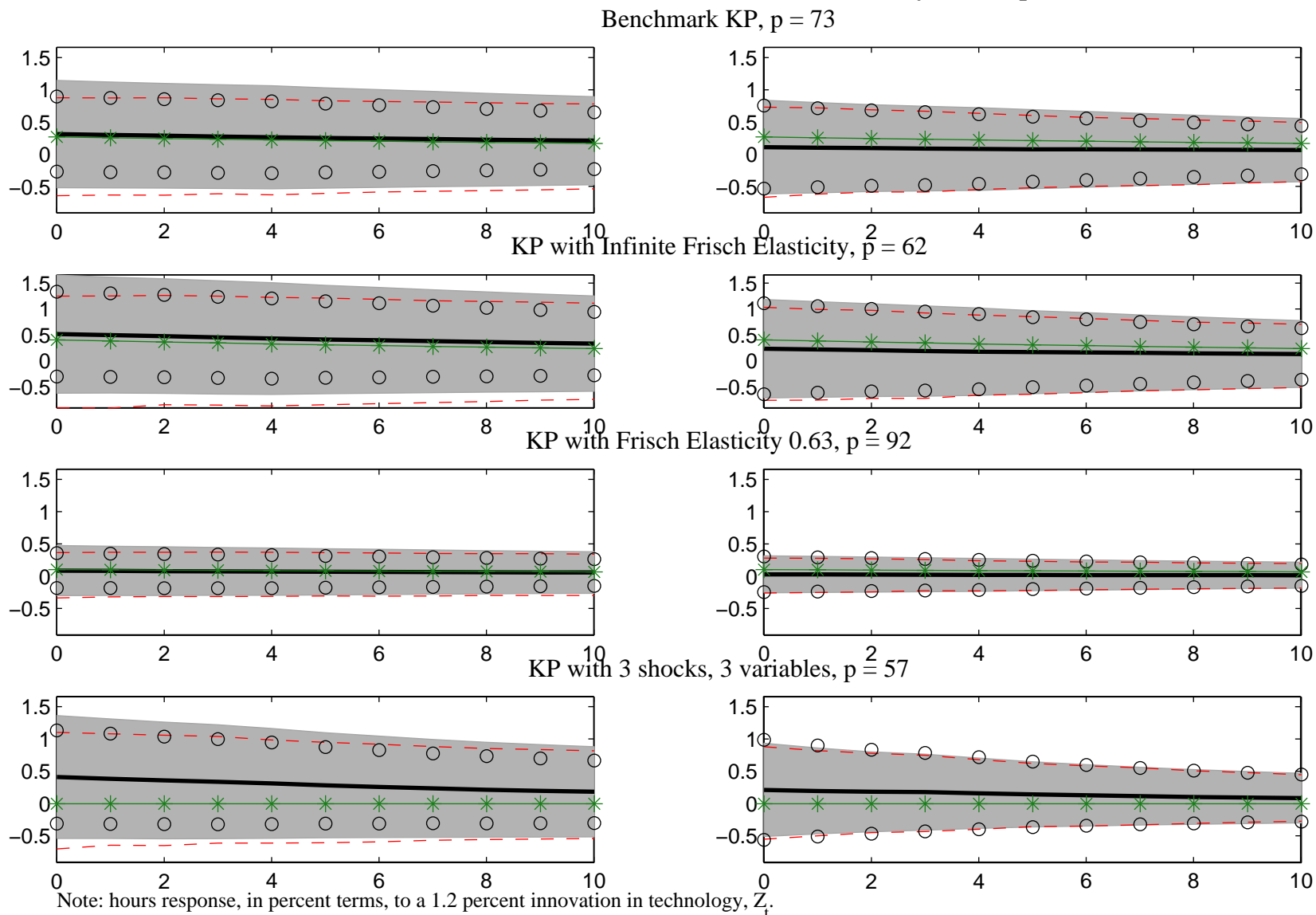
- Two key ingredients to computing impact effects of shocks:
 - Estimate of variance covariance matrix, V , of VAR disturbances and spectral density of Y_t at frequency zero, S_0 .
 - V Estimated Precisely.
 - Problem with spectral density at frequency zero.
 - * Standard VAR approach uses sum of estimated VAR matrices.
 - * No particular reason for this to be estimated precisely by ordinary least squares.
 - * Sum of lag VAR matrices corresponds to $\omega = 0$ and least squares will pay attention to this only if $S_Y(\omega)$ happens to be relatively large in a neighborhood of $\omega = 0$.
- Replace S_0 with Newey-West estimator:

$$S_0 = \sum_{k=-(T-1)}^{T-1} g(k) \hat{C}(k), \quad g(k) = \left[1 - \frac{k}{r} \right]$$

where $\hat{C}(k)$ Sample Estimate of $EY_t Y'_{t-k}$, $g(k) = 0$ for $k > r$ ($r = 150$).

- Figure 3
 - Bias is reduced
 - Less sampling uncertainty.

Figure 3: Analysis of the Long-Run Identification Assumption with Kydland-Prescott Specification

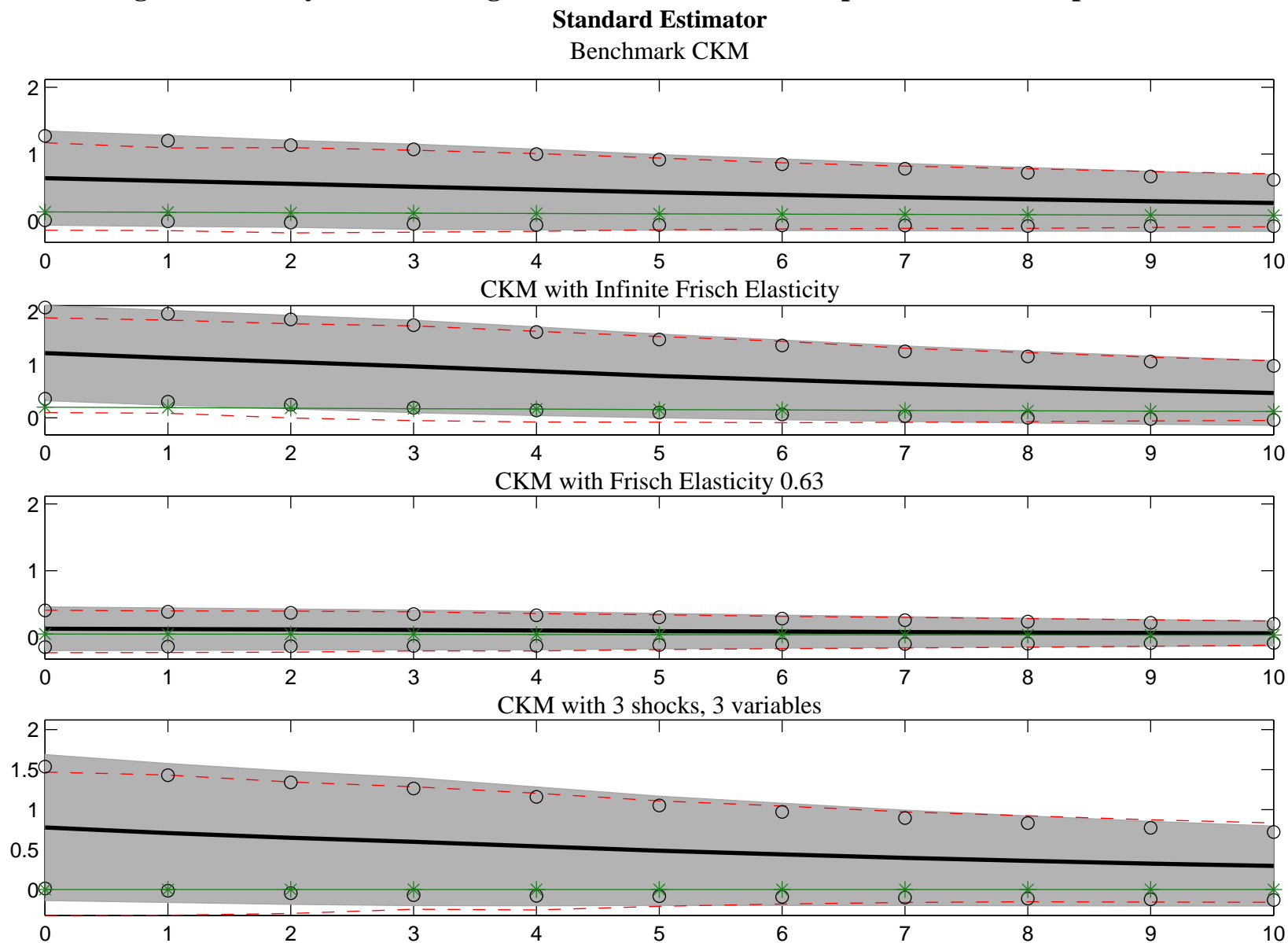


Solid line – mean response, Gray area – mean response plus/minus two standard errors,
 Starred line – true response, Dashed line – 95.5 percent probability interval of responses,
 Circles – average value of econometrician estimated plus/minus two standard errors.

CKM Long Run Results

- Benchmark CKM: substantial bias

Figure 5a: Analysis of the Long-Run Identification Assumption with CKM Specification



Note: hours response, in percent terms, to a 0.6 percent innovation in technology, Z_t .

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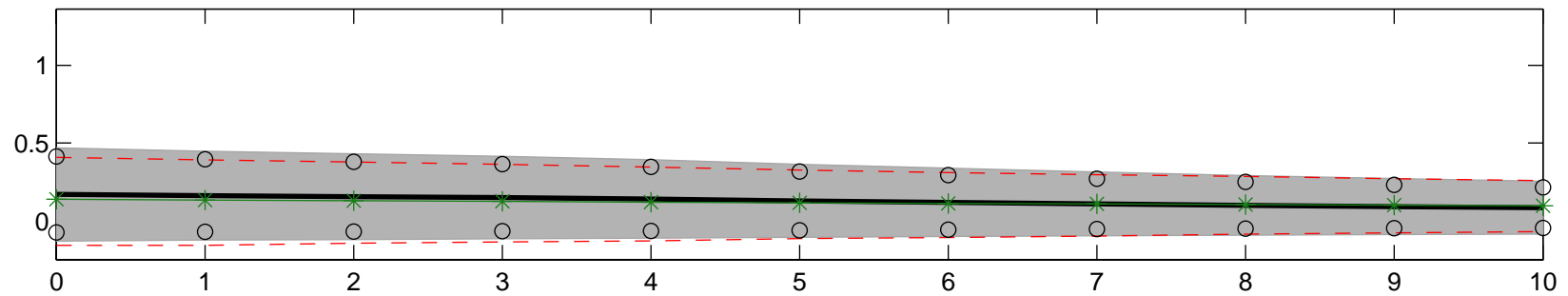
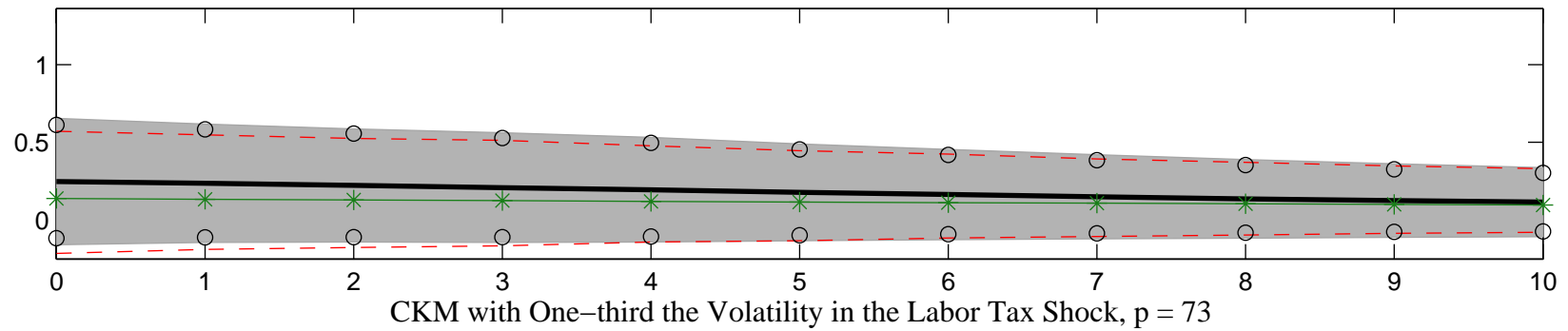
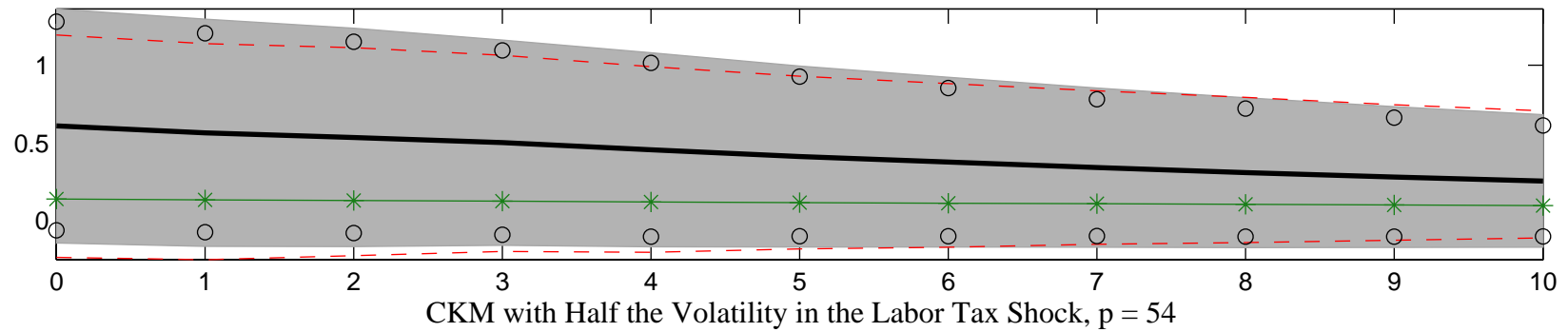
CKM Long Run Results ...

- Key Difference Between CKM and KP Model: Fraction of Variance Due to Technology Very Small (23%)

Figure 4a: Analysis of the Long-Run Identification Assumption with CKM Specification

Standard Estimator

Benchmark CKM, $p = 23$



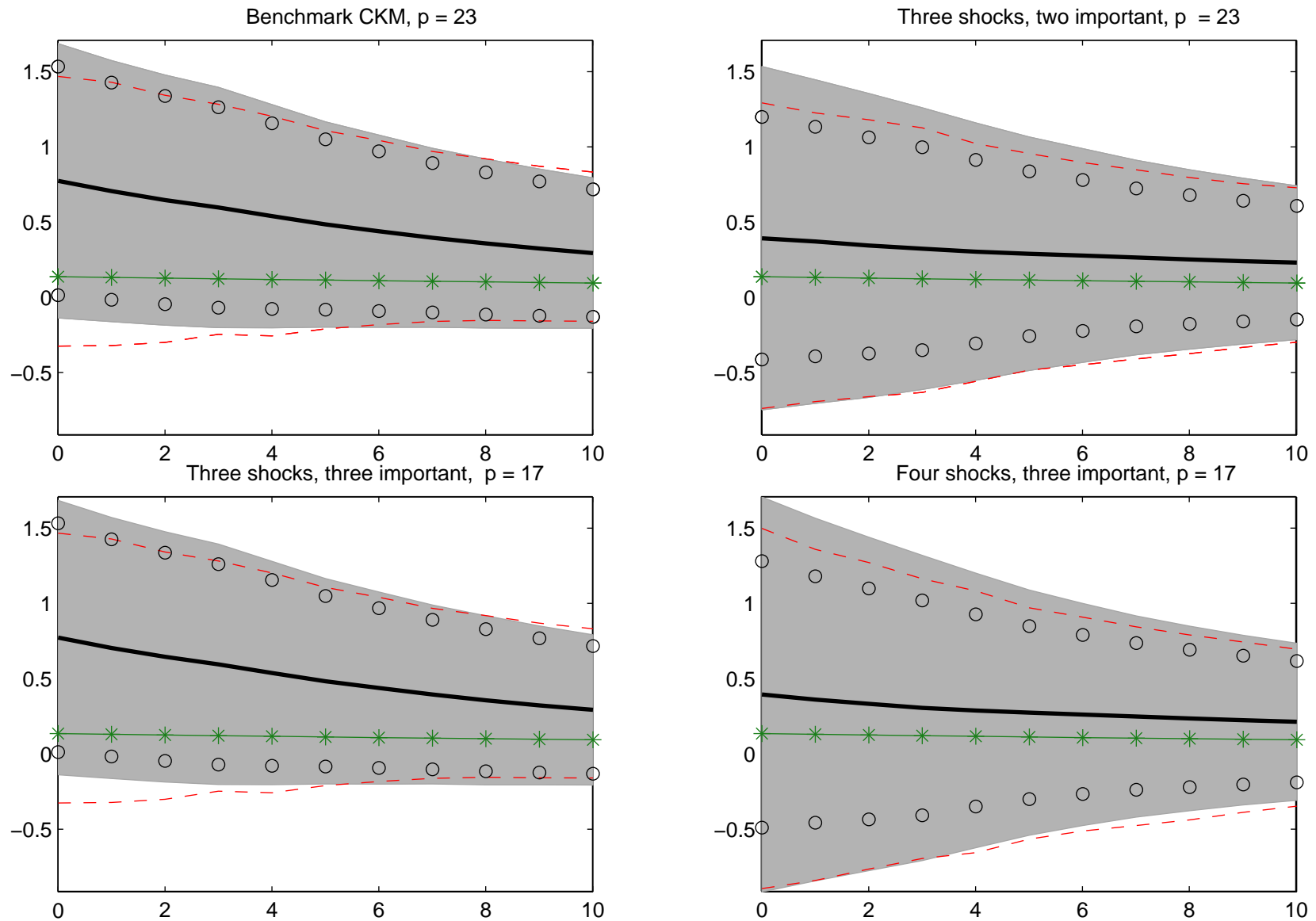
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CKM Long Run Results ...

- Distortions in CKM Model Reduced if you
 - Have One More Variable Than Important Shocks

Figure 6: The Long-Run Identification Assumption: Adding Variables and Shocks to the CKM Benchmark



Note: hours response, in percent terms, to a 0.6 percent innovation in technology, Z_t .

Estimation results for Standard VAR estimator.

Solid line – mean response, Gray area – mean response plus/minus two standard errors,

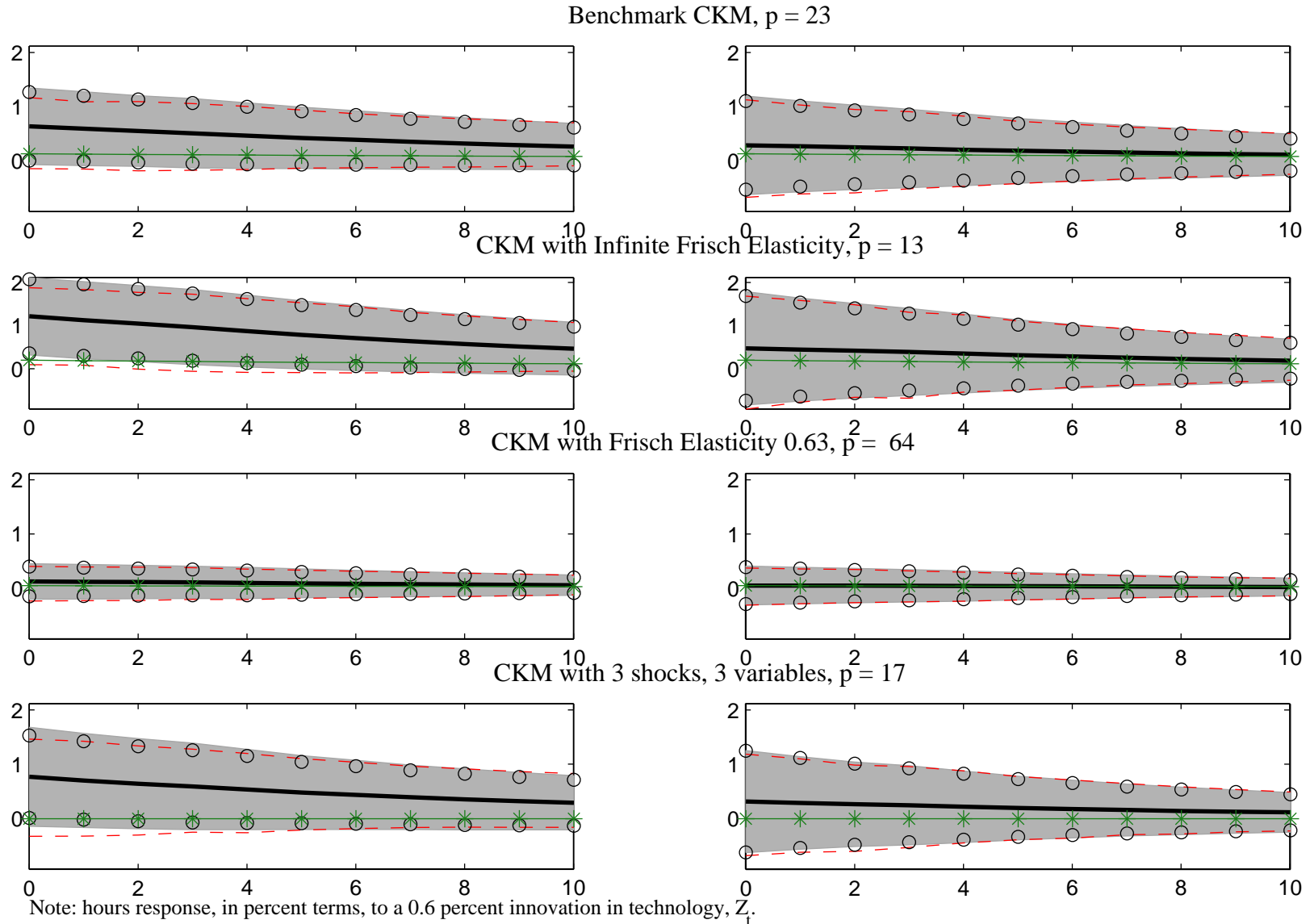
Starred line – true response, Dashed line – 95.5 percent probability interval of responses,

Circles – average value of econometrician estimated plus/minus two standard errors.

CKM Long Run Results ...

- Distortions in CKM Model Also Reduced if you
 - Adopt Newey-West Estimator of Spectrum at frequency zero.

Figure 5: Analysis of the Long-Run Identification Assumption with CKM Specification
Standard Estimator **Newey-West Spectral Estimator**



Solid line – mean response, Gray area – mean response plus/minus two standard errors,
 Starred line – true response, Dashed line – 95.5 percent probability interval of responses,
 Circles – average value of econometrician estimated plus/minus two standard errors.

Key Lessons of the RBC Model Analysis

- With Short Run Exclusion Restrictions, VAR Analysis Highly Accurate
- With Long Run Exclusion Restrictions:
 - If Technology Shocks Important, Then Inference with VARs Reliable
 - Biases Could Occur When Technology Shocks Less Important. Then,
 - * Use 5-6 Variables in VAR
 - * If Can't Use More Variables and Worried About Possibility that Technology Shocks Not Important, Use Spectral Estimator.

Why is Analysis with Short Run Restrictions So Much More Precise than with Long-Run Restrictions

- The Finding is Certainly Intuitively Appealing
 - Seems Like it Would be Tough to Find, in 50 Years' Data 'Only Shock that Has a Long-Run Effect on Productivity'
 - Shocks in Short Run Restrictions Equivalent to Regression Disturbances.

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- The VAR:

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- Short Run Restrictions:

- To Obtain Impact Effect of Shock, C_1
 - * Require Good Estimate of V
 - * That's *Exactly* What OLS Does!
- To Obtain Dynamic Effects of Shock:
 - * Require Good Estimates of B_j , first few j 's

Why is Analysis with Short Run Restrictions So Much More Precise than with Long-Run Restrictions

- Long Run Restrictions:

- To Obtain Impact Effect of Shock, C_1

- * Require Good Estimate of V and

$$B(1) = \sum_{j=1}^p B_j$$

- * OLS Provides Relatively Little Information About $B(e^{-i\omega})$, for $\omega \approx 0$.

Relationship of Our Findings to Chari-Kehoe-McGrattan

- CKM Say Nothing About Short Run Restrictions.
- CKM Consider The Consequences other Specification Errors, Such as First Differencing. We do not Consider that Here (However, see Christiano, Eichenbaum and Vigfusson, NBER Working Papers W10254 and W9819).
- CKM Overstate the Degree of Sampling Uncertainty in Estimate of Response of Hours Worked.
 - Reflects a Non-Standard Way of Implementing Long Run Restrictions

Relationship of Our Findings to Chari-Kehoe-McGrattan ...

- Impact of Shocks on Forecast of Productivity in Long-Run:

$$\lim_{j \rightarrow \infty} E_t[a_{t+j}] - E_{t-1}[a_{t+j}] = (1, 0) [I - B(1)]^{-1} C_1 e_{1t}$$

- Standard Implementation of Long Run Restrictions:

(1, 1) Element of $[I - B(1)]^{-1} C_1$ Must Be Positive (Long Run Effect)
Sign of C_{11} (Impact Effect of Technology Shock) unrestricted
Could Lead to Contemporaneous Drop in Productivity

- CKM Sign Restriction:

CKM Sign Restriction: $C_{11} > 0$,

(1, 1) Element of $[I - B(1)]^{-1} C$ unrestricted

‘Positive Technology Shock Leads to Contemporaneous Rise in Productivity’

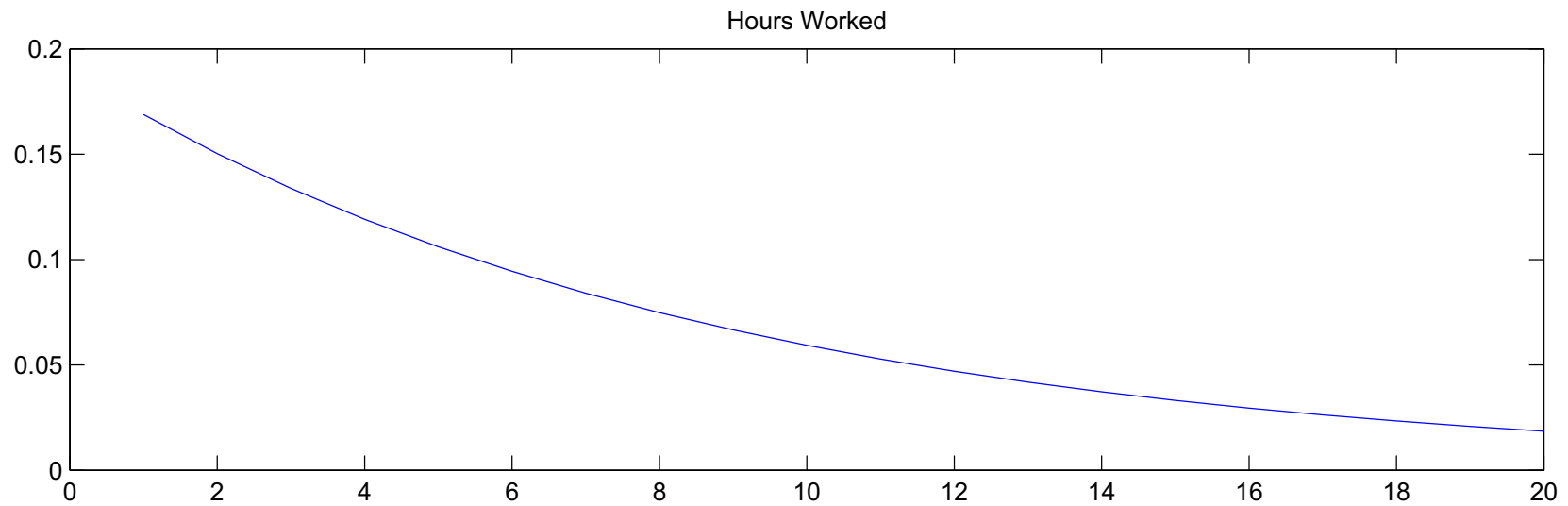
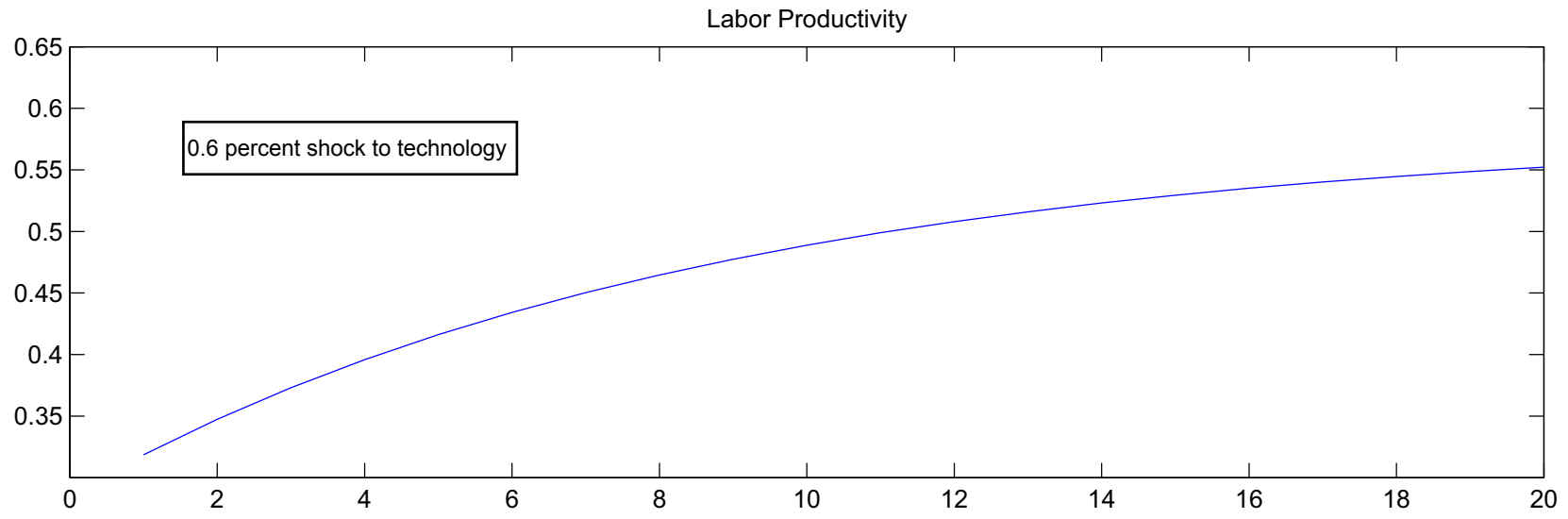
Positive Technology Shock Could Lead to Permanent *Reduction* in Productivity

(This Pattern is Impossible in CKM DGP)

Relationship of Our Findings to Chari-Kehoe-McGrattan ...

- Benchmark CKM Model
 - Initial Hours Worked Response Strong, Productivity Response Weak (Fig 3)

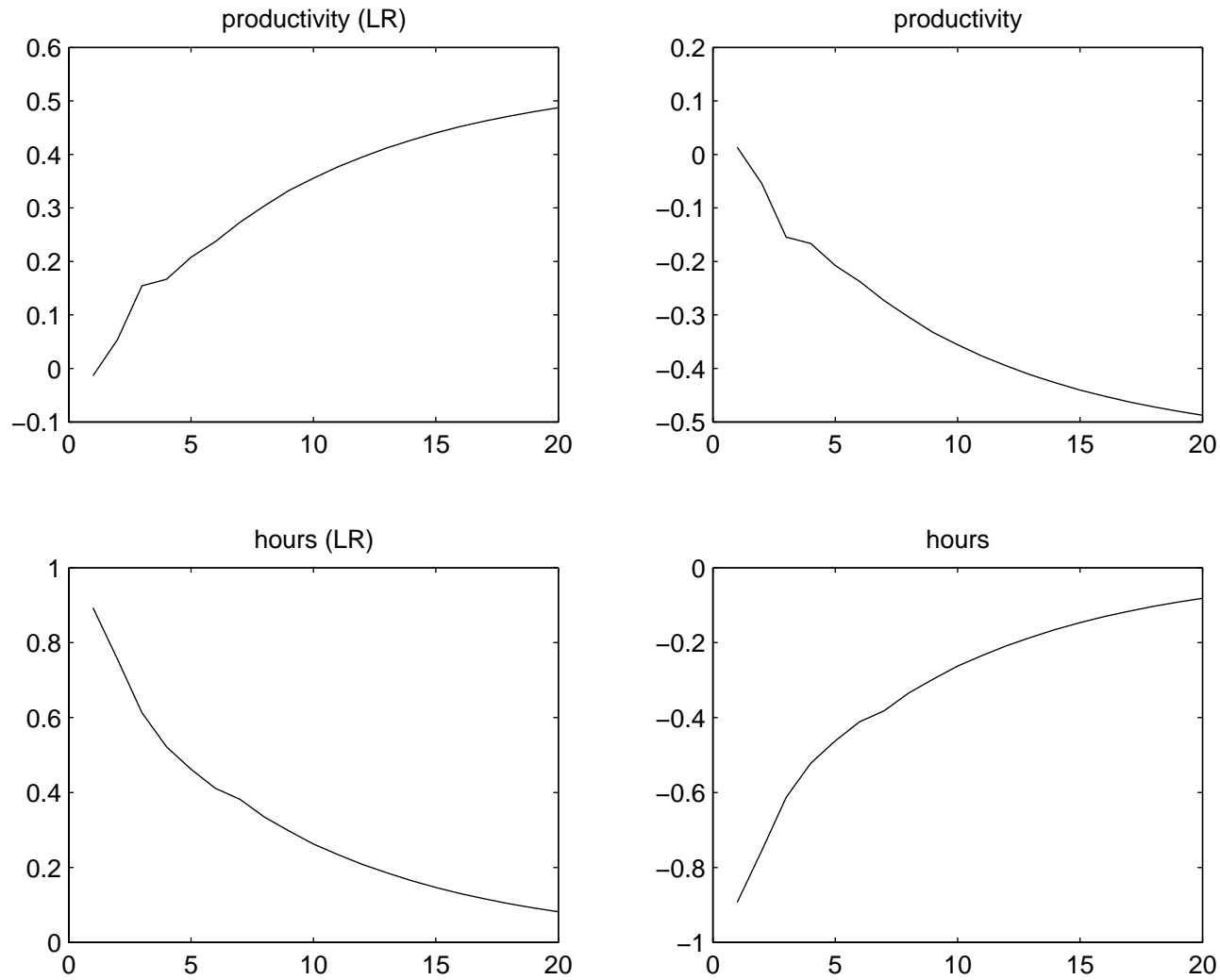
Response of Hours Worked and Labor Productivity in Benchmark CKM Example



Relationship of Our Findings to Chari-Kehoe-McGrattan ...

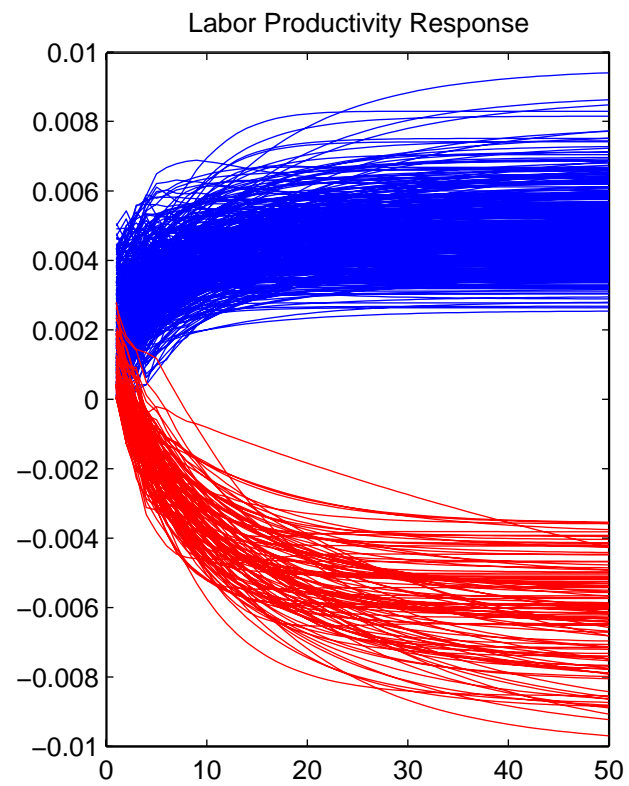
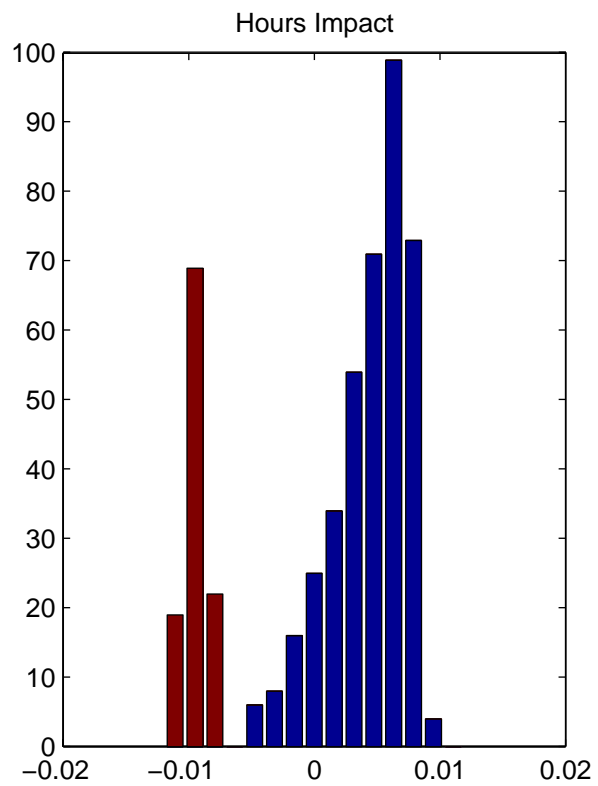
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Figure 3a: Example of Mistaken Inference



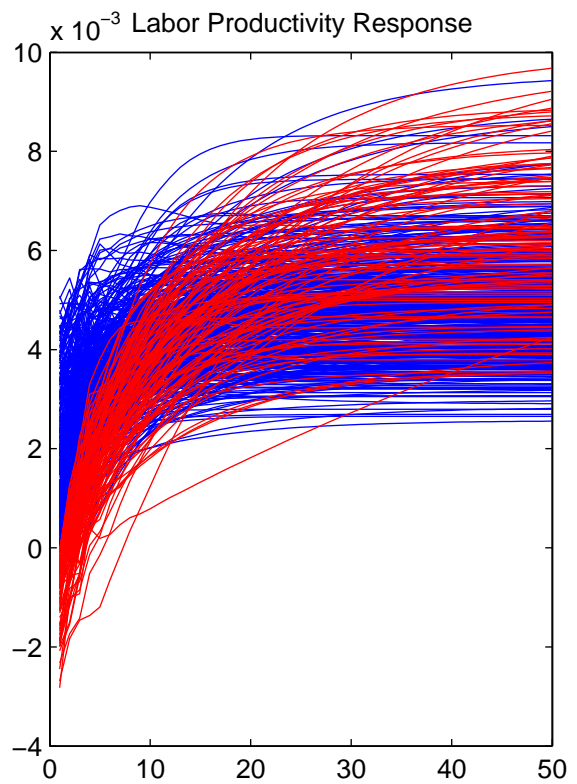
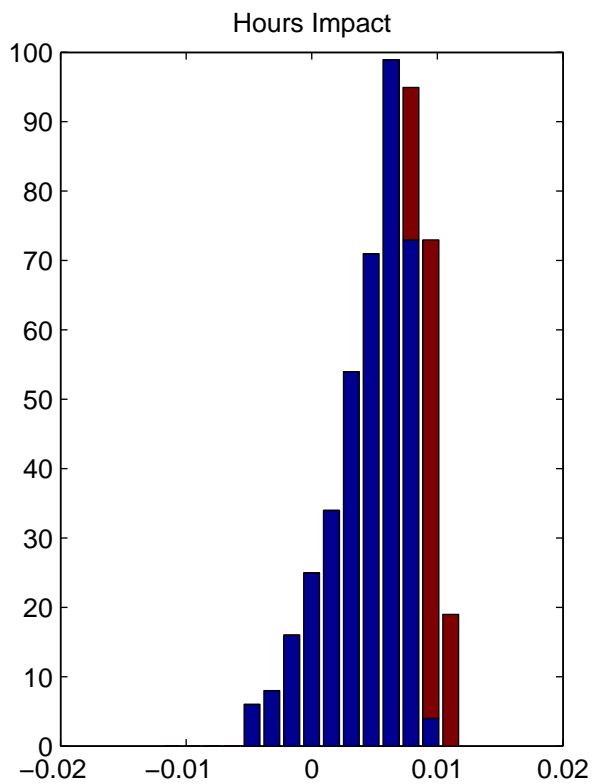
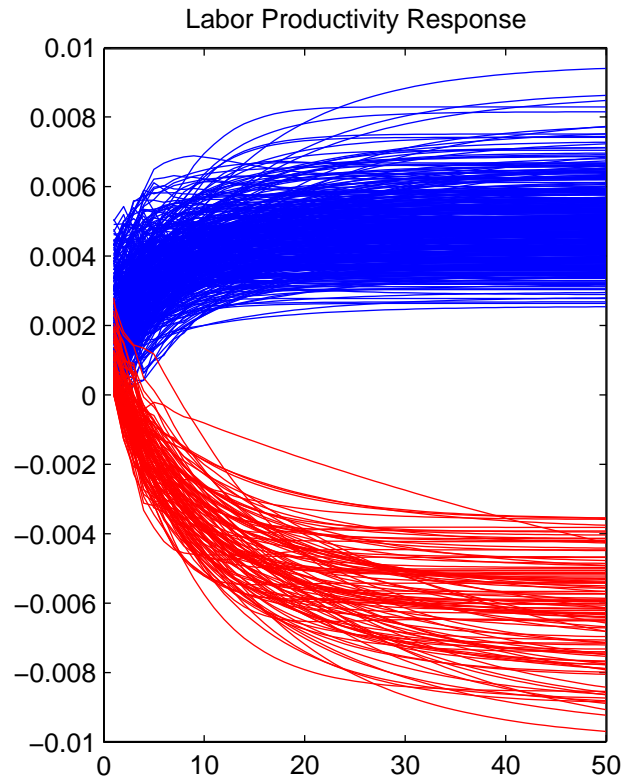
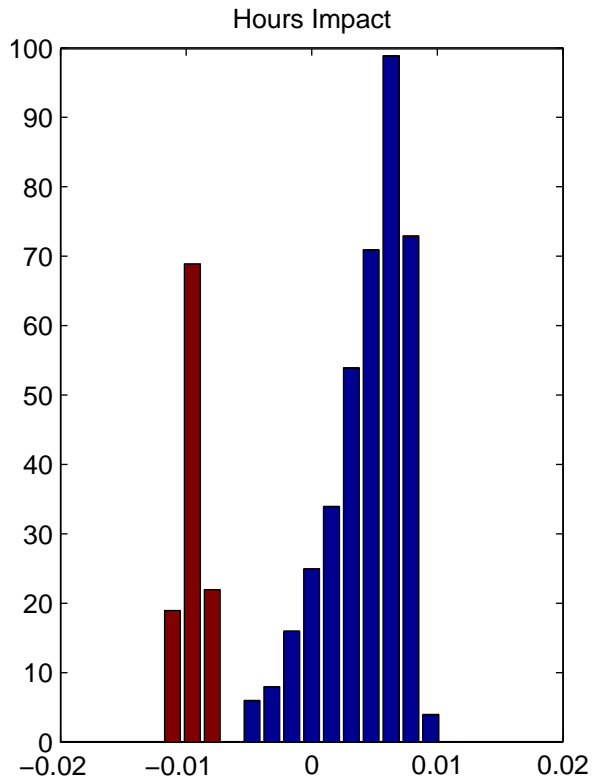
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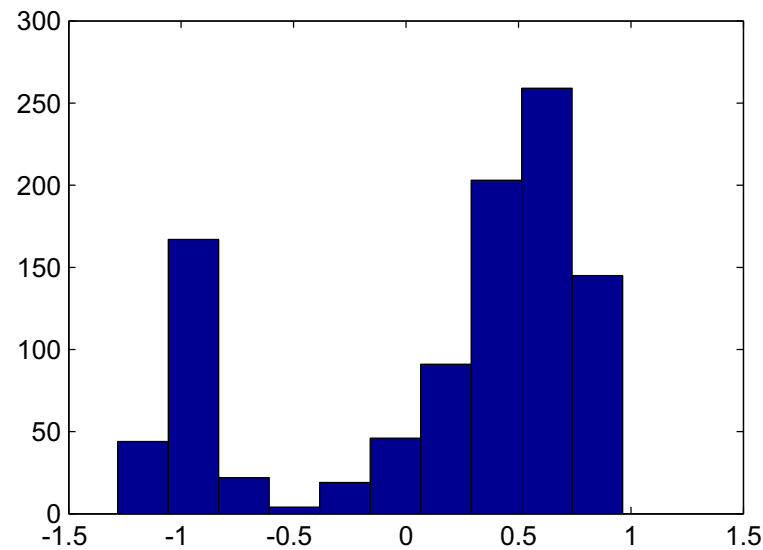
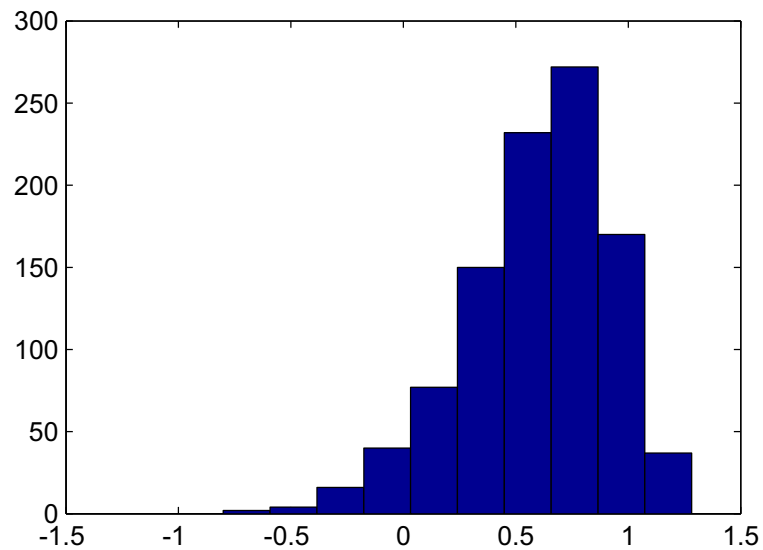
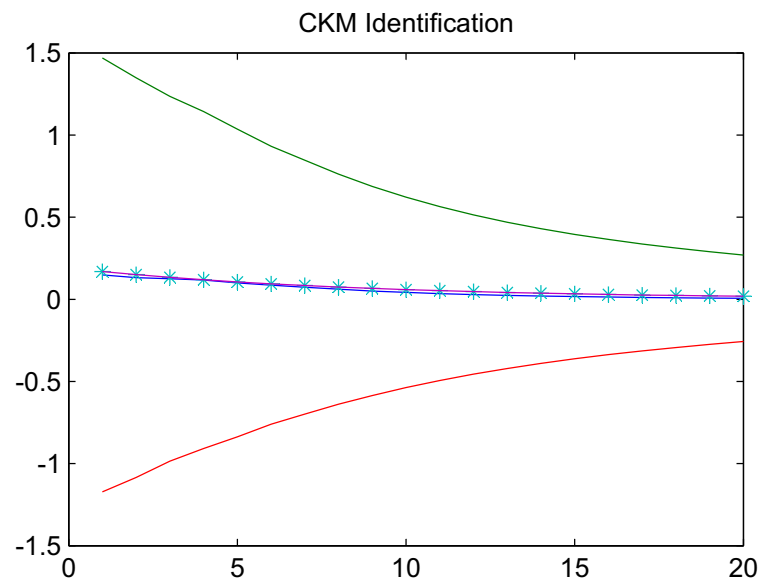
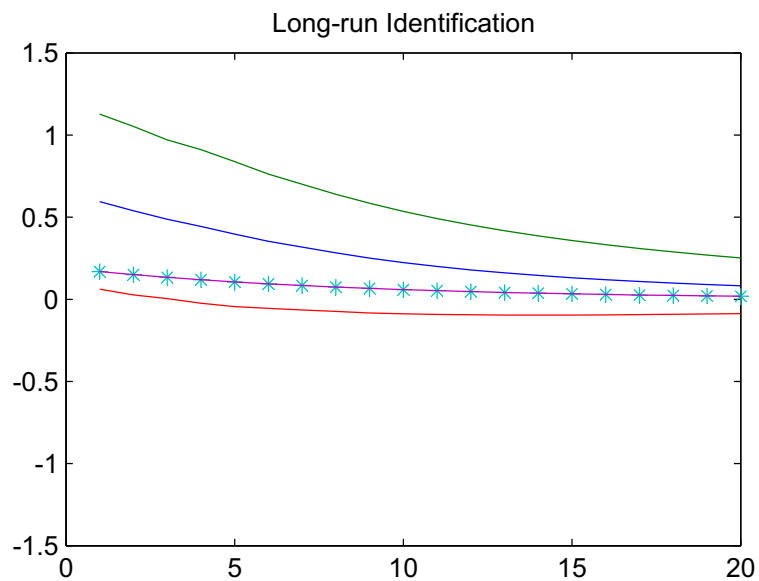
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 - How Does this Affect Standard Errors?
 - * Leads CKM to Substantially Overstate Sampling Uncertainty (Fig. 5)

Impact of CKM Sign Restriction on Sampling Uncertainty



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 - Potentially perform less well when shock under investigation isn't very important.
 - When there are enough variables in VAR, problems are greatly mitigated.
- Develop and implement a modified VAR approach
 - Leads to drastic improvement even when technology shocks play a limited role in aggregate fluctuations and a small number of variables are included in VAR.