Online Appendix for "The Baby Boom and World War II: A Macroeconomic Analysis"

Matthias Doepke Moshe Hazan Yishay D. Maoz

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A International Evidence: Allied versus Neutral Countries in Terms of Period Fertility

In the main text, we use cohort fertility rates to compare the baby-boom experience of two groups of countries: the Allied countries that, like the United States, did not fight on their own soil (Australia, Canada, and New Zealand), and the major European countries that remained neutral in the war (Ireland, Portugal, Spain, Sweden, and Switzerland). Here we show that our results are robust to instead doing the comparison in terms of period fertility rates, measured by the total fertility rate. Figures A1 and A2 compare the evolution of the total fertility rate in the United States to that in the Allies countries (Figure A1) and to that of neutral European countries (Figure A2). Just as in the data for completed fertility displayed in the paper, the Allied countries experience baby booms remarkably similar to the U.S. case in terms of magnitude and timing. In contrast, none of the neutral countries experiences a baby boom on the same scale. In Sweden and Switzerland, there is some rise in fertility in the early 1940s (in the aftermath of the Great Depression, which we will further discuss below), but there is no rise in fertility from 1945 to the 1950s, which is when the large baby boom is observed in the Allied countries, and when our mechanism should be at work.

The only country from our comparison group for which total fertility rates are not available for this period is Ireland. In terms of completed fertility (displayed in the paper), Ireland looks just like the other neutral countries, albeit with a higher level of fertility. To provide an impression of period fertility in Ireland compared to the United States, Figure A3 displays the birth rate data for Ireland and the United States from 1930 to 1970. The picture is fully consistent with the other evidence already discussed. Both countries experience a drop in fertility during the Great Depression followed by a recovery until the early 1940s. After World War II, there is a further takeoff in fertility in the United States towards the peak of the baby boom, but a moderate decline in Ireland.

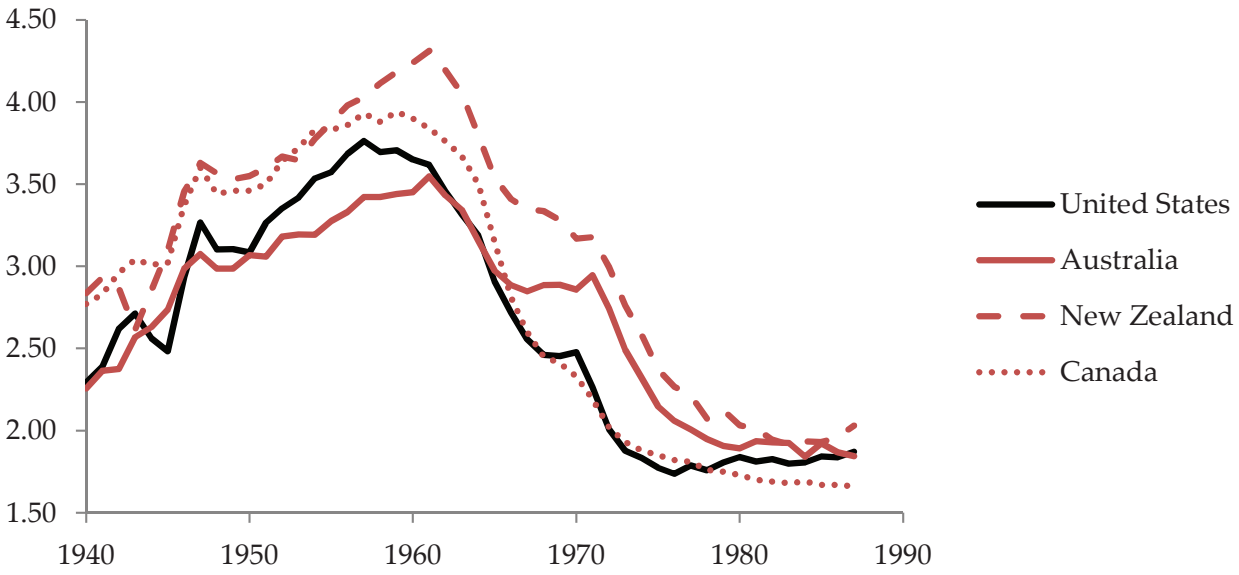

Figure A1: Total Fertility Rate in United States and Allied Countries (Source: Observatoire Démographique Européen)

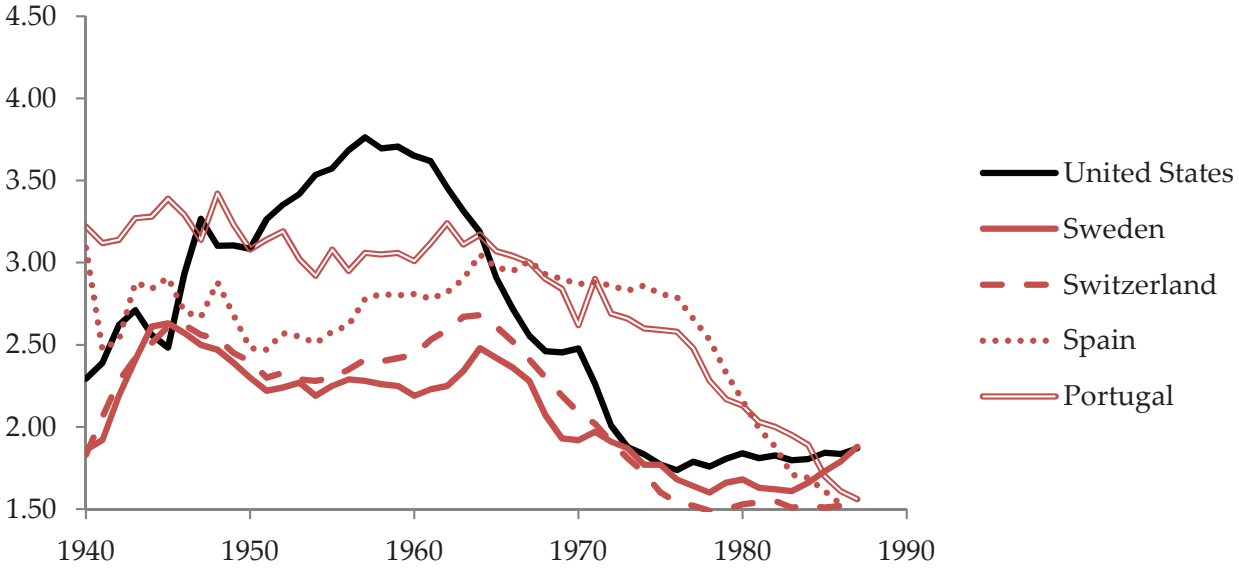

Figure A2: Total Fertility Rate in United States and Neutral Countries (Source: Observatoire Démographique Européen)

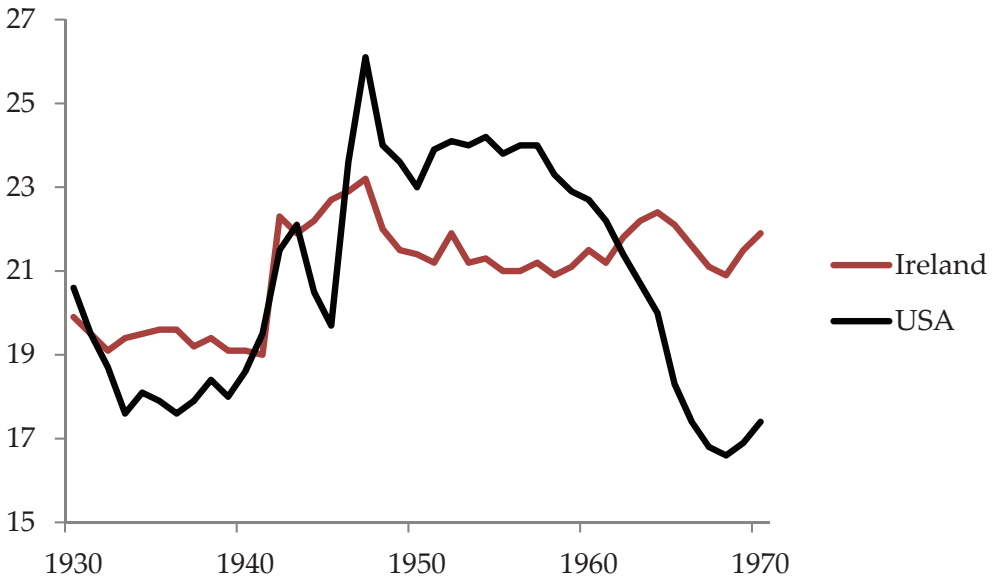

Figure A3: Birth Rates in Ireland and the United States (Source: Mitchell, Brian R. International Historical Statistics: The Americas , 1750–1993. New York, N.Y.: Stockton Press, 1998, and: Mitchell, Brian R. International Historical Statistics: Europe, 1750–1993. New York, N.Y.: Stockton Press, 1998.

B Additional Regression Results

In this section, we display a number of additional regression results as a check of the robustness of our main findings.

B.1 Additional State-Level Controls

One concern about the mobilization variable is that mobilization is not randomly assigned. Acemoglu, Autor, and Lyle (2004) show that the mobilization rate is significantly related to a number of state-level variables as of 1940, the most important ones being the share of the age-relevant 1940 male population in farm occupations, the share of the male population that is black, and the share of German ancestry. It is possible that factors that led to cross-state differences in mobilization also had effects on changes in fertility, marriage, and labor supply from 1940 to 1960 through other channels. For example, if there were differences in general economic trends between more and less agricultural states, this may also affect fertility and labor supply. The main concerns in our case are the role of the share of agriculture and the share of the black population. We partly address such concerns by including measures of race and agricultural employment at the individual level in all our regressions, and letting the coefficients vary by year, which should capture general changes in the impact of agricultural employment and race on fertility, employment, and marriage. Another approach, which we pursue here, is to directly include known state-level determinants of the mobilization rate in the main regressions (results of this kind are also presented in Acemoglu, Autor,

and Lyle 2004). Such regressions yield the impact of the "unexplained" variation in the mobilization rate. Even with this approach there may be omitted variable bias, because the unexplained variation is not randomly assigned either. Also, some of the variation due to the other factors may in fact be related to our mechanism; for example, agricultural states were presumably less likely to have many manufacturing establishments that were relevant for wartime production, and hence there may be an effect on the female labor market that goes beyond the fact that some men had to stay to provide farm labor. Nevertheless, including these additional regressors can be a useful diagnostic for isolating potential concerns about the identification of the results.

The additional state-level determinants of mobilization that we consider are the fraction of men aged 13–44 who were born in Germany, the fraction of men aged 13–44 who are black, and the fraction of men aged 13–44 who have farm occupations (all in 1940). Table A1 displays results for fertility, labor supply, and marriage of young women aged 25–35 with these additional controls included. As the baseline specification, we employ what is displayed in column (2) of Table 2, i.e., our preferred specification with assignment of mobility by state of birth and including year-specific controls for agricultural employment and years of education. Table A2 presents analogous results for the employment of older women aged 45–55. In each case, the new controls are interacted with the 1960 indicator variable.[1] What is displayed is the coefficient estimate for mobilization interacted with 1960 (our variable of interest) and the coefficient estimates for the additional controls interacted with 1960.

The results can be summarized as follows. Adding a control for the German share reduces the size of the effect of mobilization in some cases, but all previously significant results stay significant and remain quantitatively large. The same applies when the share of the age-relevant male black population is also controlled for, except that the effect on the probability of employment for the old now becomes insignificant. Finally, when the age-relevant share of male employment in agriculture is also controlled for, the estimates for children under age 5 and marriage are further reduced in size, but remain significant at the 10 percent level, and still imply a quantitatively important effect of mobilization on fertility.[2] In contrast, the estimates for the probability of employment and weeks worked retain the same sign (negative for the young and positive for the old), but are no longer statistically significant.

To some extent, the reduced significance of the estimated employment effects is not surprising. Acemoglu, Autor, and Lyle (2004) also found smaller employment effects with additional controls in 1950, and by 1960 we would expect the employment effects of the

[1] Adding the controls in levels has no effect, because state-level dummy variables were already included.

[2] Moreover, if we run the fertility regression in a non-linear specification the precision is further improved; the coefficient for the outcome children under age 5 is significant at the 5 percent level in an ordered probit specification, and at the 1 percent level in a negative binomial specification. In a linear regression with a discrete fertility variable (1 if a child under the age of 5 is present) as reported in Table A5 below, including all three controls again leads to a coefficient estimate that is significant at the 5 percent level.

Table A1: Impact of WWII Mobilization Rates on Fertility, Labor Supply, and Marriage of Women Aged 25–35 with Additional State-Level Controls

Independent Variable	Dependent Variable			
	Children Under Age 5			
Mobilization × 1960	0.953 (0.237)	0.815 (0.250)	0.547 (0.265)	0.667 (0.382)
Share German × 1960		5.041 (1.495)	4.136 (1.370)	4.575 (1.734)
Share Black × 1960			-0.150 (0.076)	-0.151 (0.073)
Share Farming × 1960				0.056 (0.114)
	Children Ever Born			
Mobilization × 1960	0.818 (0.543)	0.634 (0.496)	-0.715 (0.445)	0.551 (0.547)
Share German × 1960		6.202 (3.291)	1.528 (2.424)	6.248 (2.716)
Share Black × 1960			-0.745 (0.155)	0.757 (0.155)
Share Farming × 1960				0.598 (0.180)
	Employed			
Mobilization × 1960	-0.608 (0.137)	-0.481 (0.154)	-0.561 (0.168)	-0.165 (0.151)
Share German × 1960		-4.521 (0.555)	-4.794 (0.529)	-3.345 (0.531)
Share Black × 1960			-0.045 (0.037)	-0.049 (0.032)
Share Farming × 1960				0.185 (0.028)
	Weeks Worked			
Mobilization × 1960	-23.173 (7.065)	-17.696 (7.807)	-19.684 (8.615)	-5.803 (8.077)
Share German × 1960		-193.99 (28.0)	-200.7 (27.4)	-149.86 (28.89)
Share Black × 1960			-1.112 (1.506)	-1.240 (1.471)
Share Farming × 1960				6.487 (1.622)
	Ever Married			
Mobilization × 1960	0.488 (0.119)	0.433 (0.140)	0.384 (0.152)	0.252 (0.146)
Share German × 1960		1.951 (0.450)	1.785 (0.451)	1.302 (0.506)
Share Black × 1960			-0.028 (0.021)	-0.026 (0.022)
Share Farming × 1960				-0.062 (0.034)

Notes: Standard errors (in parentheses) are adjusted for clusters of state of birth and year of observation. Estimates are from regressions of pooled micro data from the 1940 and 1960 censuses. All regressions are OLS. All regressions include indicator variables of observation year, age, race, state/country of birth, years of education, and farm status. Regressions for children under age 5 also contain number of children older than 5. Where applicable, the number of older children and the farm, race, and age indicator variables are also interacted with the 1960 indicator variable. For children ever born, the sample is restricted to observation where the number of children does not exceed seven. The number of observations is 132,934 for children ever born and 225,613 for the other variables. All data are weighted using census person weights (sample line weights for children ever born).

Table A2: Impact of WWII Mobilization Rates Labor Supply of Women Aged 45–55 with Additional State-Level Controls

Independent Variable	Dependent Variable			
	Employed			
Mobilization × 1960	0.231 (0.084)	0.173 (0.077)	-0.014 (0.098)	0.009 (0.136)
Share German × 1960		2.113 (0.798)	1.434 (0.715)	1.514 (0.769)
Share Black × 1960			-0.110 (0.031)	-0.110 (0.032)
Share Farming × 1960				0.010 (0.038)
	Weeks Worked			
Mobilization × 1960	18.390 (3.775)	15.334 (3.248)	8.325 (3.525)	6.272 (5.198)
Share German × 1960		111.99 (30.17)	86.475 (27.041)	79.31 (29.902)
Share Black × 1960			-4.136 (1.264)	-4.125 (1.257)
Share Farming × 1960				-0.933 (1.645)

Notes: Standard errors (in parentheses) are adjusted for clusters of state of birth and year of observation. Estimates are from regressions of pooled micro data from the 1940 and 1960 censuses. All regressions are OLS. All regressions include indicator variables of observation year, age, race, state/country of birth, years of education, and farm status. The farm, race, and age indicator variables are also interacted with the 1960 indicator variable. The number of observations is 164,408. All data are weighted using census person weights.

Table A3: Impact of WWII Mobilization Rates on Fertility, Labor Supply, and Marriage of Women Aged 25–35 and Labor Supply of Women 45–55 for White Non-Farm Sample (Coefficient Estimates from OLS Regressions for Variable "Mobilization Rate $\times$ 1960")

Dependent Variable	Age Group:					
	Age 25–35 (N = 174,648)			Age 45–55 (N = 125,614)		
	(1)	(2)	(3)	(4)	(5)	(6)
Children Under Age 5	1.096	1.040	0.960			
	(0.250)	(0.240)	(0.211)			
R^2	0.116	0.118	0.204			
Children Ever Born	0.439	0.965	0.993			
	(0.693)	(0.570)	(0.567)			
R^2	0.108	0.148	0.154			
Employed	-0.544	-0.635	-0.369	0.215	0.196	0.244
	(0.168)	(0.157)	(0.103)	(0.086)	(0.079)	(0.067)
R^2	0.007	0.051	0.210	0.059	0.086	0.189
Weeks Worked	-21.746	-26.412	-13.889	14.573	14.317	16.473
	(8.520)	(7.779)	(5.054)	(3.978)	(3.503)	(3.569)
R^2	0.009	0.025	0.223	0.050	0.072	0.185
Ever Married	0.433	0.519				
	(0.140)	(0.139)				
R^2	0.050	0.071				
Education Controls	no	yes	yes	no	yes	yes
Marital Status Controls	no	no	yes	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of birth and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. All regressions are OLS. Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state/country of birth. Regressions for children under Age 5 also contain number of children older than 5. Where applicable, the number of older children and the race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rate assigned by state of birth. For Children Ever Born, the sample is restricted to observation where the number of children does not exceed seven. All data are weighted using census person weights (sample line weights for children ever born). The number of observations for the outcome "children ever born" is 107,873.

war to be further muted, so that using only the residual variation in mobilization should result in even noisier estimates.[3] Multicollinearity is also a concern. For example, in the weeks worked regression for women 45–55 introducing the farming share renders the coefficient on mobilization insignificant, but the share on farming itself is also insignificant. An F-test reveals that the two coefficients are jointly significant with a p-value of 0.053. Thus, correlation among the explanatory variables may make it difficult to distinguish effects that are due to mobilization itself as opposed to the determinants of mobilization directly.

In summary, directly including state-level determinants of mobilization further strengthens the evidence of an important impact of mobilization on the change in fertility and marriage from before the war to the peak of the baby boom. While lacking randomized mobilization we cannot fully rule out omitted variable bias, it is also useful to notice that to explain the facts, a potential omitted variable would not just have to result in differential trends between high- and low-mobilization states, but this differential trend would also need to have opposite effects on the employment of young versus old women. It is difficult to come up with a mechanism that has this feature other than the labor-market mechanism that we develop here.

The results do raise some concerns about the roles of agriculture and race for the change in employment of young and old women. As a further robustness check to deal with this concern (in addition to including year-dependent race and agricultural employment dummies in our main regressions), we also repeated our main regressions with a sample restricted to white women who are not employed on a farm. Table A3 displays results analogous to columns (1) to (3) of Tables 2 and 4 for the white non-farm sample. These results turn out to be even stronger than our baseline results. Once education is controlled for, for the young group the estimated effects on fertility, employment, and marriage are all larger than in the general sample. Unlike in Table 2, the estimated effect on the number of children ever born is now significant at the 10 percent level. For the 45–55 age group, the effects on the probability of employment and weeks worked are now significant in all three specifications. These findings underline that our main results are not driven by changes that are conditional on race or agriculture.

B.2 Allowing for Nonlinear Effects

Next, we confront the issue that the linear regression model may not be ideally suited for all the variables that we consider. Specifically, employed and ever married are binary variables, children ever born is a discrete variable, and children under age 5 is a discrete variable with a limited range. To check the robustness of our findings to alternative regression models, Table A4 displays results analogous to those in Tables 2 and 4 using ordered probit (children under age 5 and children ever born) and probit (employed and ever married) regression. The results are qualitatively and quantitatively similar to

[3]Note that unlike employment, the fertility and marriage variables reflect decisions taken over a longer period, and thus should be influenced by employment conditions over a number of years.

Table A4: Impact of WWII Mobilization Rates on Fertility, Female Labor Supply, and Probability of Marriage (Coefficient Estimates from Probit and Ordered Probit Regressions for Variable "Mobilization Rate $\times$ 1960")

Dependent variable	Regression		
	(1)	(2)	(3)
	Age 25-35 (N = 225,613)		
Children Under Age 5	1.982	1.232	0.992
	(0.302)	(0.274)	(0.236)
Children Ever Born	1.537	0.666	0.673
	(0.494)	(0.391)	(0.391)
Employed	-2.702	-1.754	-1.083
	(0.442)	(0.395)	(0.353)
Ever Married	0.769	0.887	
	(0.413)	(0.416)	
	Age 45-55 (N = 164,408)		
Employed	-0.238	0.532	0.788
	(0.307)	(0.327)	(0.321)
Education and Farm Controls	no	yes	yes
Marital Status Controls	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of birth and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. Regressions are probit (Employed, Ever Married) and ordered probit (children under age 5, children ever born). Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, state of residence, and state/country of birth. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rates are assigned by state of birth. All data are weighted using census person weights (sample line weights for children ever born). For children ever born, the sample is restricted to observation where the number of children does not exceed seven. Number of observations for children ever born is 132,934.

Table A5: Impact of WWII Mobilization Rates on Fertility for Discrete Fertility Measure (Coefficient Estimates from OLS Regressions of Indicator Variable for "At Least One Child Under Age 5" for Variable "Mobilization Rate $\times$ 1960")

Dependent variable	Regression			
	(1)	(2)	(3)	(4)
	Age 25-35 (N = 225,613)			
At Least One Child Under Age 5	0.893	0.554	0.417	0.389
	(0.134)	(0.112)	(0.081)	(0.172)
Education and Farm Controls	no	yes	yes	yes
Marital Status Controls	no	no	yes	no
Additional State-Level Controls	no	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of birth and year of observation. Estimates are from separate OLS regressions of pooled micro data from the 1940 and 1960 censuses. The outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state/country of birth. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. The additional controls are the age-relevant share of the male population that is German, black, or in farming occupations, in each case interacted with the 1960 indicator variable. All data are weighted using census person weights.

the linear results displayed in the main text. The most important change is that with the nonlinear specification, the parameter estimates for children ever born are now all significant (with p-values of 0.002, 0.08, and 0.085 in columns (1) to (3)).

Another approach to deal with nonlinearity in the fertility measure of children under age 5 is to collapse the measure to a binary variable that is one if at least one child under 5 is present, and zero if not. Table A5 displays results for this variable. The coefficient estimates for all three specifications are large and precisely estimated. In column (4) we add the three additional state-level controls that we used in Tables A1 and A2 to our preferred specification. Even with the additional controls, the coefficient estimate remains large and statistically significant at the 5 percent level.

B.3 State of Birth Controls and Controlling for Mobility

Another possible concern is that our regressions that assign mobility by state of residence control for both state-of-birth and state-of-residence effects. Tables A6 and A7 display results for our main regressions with state-of-birth effects omitted with our baseline results that include these effects. As can be seen, the new results are virtually identical to the baseline.

Next, we consider the issue that we do not directly observe a woman's state of residence when she was affected by the mobilization shock, but we have to use either the state of

Table A6: Impact of WWII Mobilization Rates on Fertility, Labor Supply, and Marriage of Women Aged 25–35 (Coefficient Estimates from OLS Regressions for Variable "Mobilization Rate $\times$ 1960")

Dependent Variable	Regression					
	State of Birth Controls			No State of Birth Controls		
	(1)	(2)	(3)	(4)	(5)	(6)
Children Under Age 5	1.146	0.665	0.573	1.12	0.641	0.541
	(0.259)	(0.232)	(0.208)	(0.256)	(0.23)	(0.205)
R^2	0.115	0.119	0.196	0.113	0.118	0.194
Children Ever Born	2.425	0.781	0.749	2.341	0.720	0.694
	(0.741)	(0.540)	(0.537)	(0.729)	(0.524)	(0.521)
R^2	0.092	0.142	0.147	0.089	0.138	0.143
Employed	-0.820	-0.395	-0.204	-0.821	-0.395	-0.202
	(0.149)	(0.129)	(0.096)	(0.15)	(0.131)	(0.095)
R^2	0.020	0.046	0.204	0.018	0.045	0.203
Weeks Worked	-26.022	-12.338	-3.558	-26.253	-12.528	-3.619
	(7.670)	(7.087)	(5.066)	(7.754)	(7.205)	(5.074)
R^2	0.020	0.041	0.203	0.018	0.039	0.202
Ever Married	0.384	0.377		0.397	0.388	
	(0.119)	(0.122)		(0.121)	(0.124)	
R^2	0.046	0.063		0.043	0.061	
Educ. and Farm Controls	no	yes	yes	no	yes	yes
Marital Status Controls	no	no	yes	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of residence and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state of residence. Regressions for children under age 5 also contain number of children older than 5. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rates are assigned by state of residence. All data are weighted using census person weights (sample line weights for children ever born). For children ever born, the sample is restricted to observation where the number of children does not exceed seven. The number of observations is 243,554 (children ever born: 143,637).

Table A7: Impact of WWII Mobilization Rates on Labor Supply of Women Aged 45–55 (Coefficient Estimates from OLS Regressions for Variable "Mobilization Rate × 1960")

Dependent Variable	Regression					
	State of Birth Controls			No State of Birth Controls		
	(1)	(2)	(3)	(4)	(5)	(6)
	Age 45–55 (N = 191,715)					
Employed	0.058	0.185	0.174	0.133	0.195	0.134
	(0.074)	(0.081)	(0.077)	(0.08)	(0.084)	(0.076)
R^2	0.081	0.112	0.200	0.079	0.111	0.199
Weeks Worked	17.998	17.349	16.552	21.646	17.838	14.697
	(4.384)	(4.326)	(4.641)	(4.575)	(4.329)	(4.414)
R^2	0.066	0.088	0.182	0.063	0.087	0.181
Educ. and Farm Controls	no	yes	yes	no	yes	yes
Marital Status Controls	no	no	yes	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of residence and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state of residence. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rates are assigned by state of residence. All data are weighted using census person weights.

Table A8: Impact of WWII Mobilization Rates on Fertility, Labor Supply, and Marriage of Women Aged 25–35 (Coefficient Estimates from OLS Regressions for Variable "Mobilization Rate $\times$ 1960")

Dependent Variable	Regression					
	Full Sample			Stayers Sample		
	(1)	(2)	(3)	(4)	(5)	(6)
Children Under Age 5	1.146	0.665	0.573	1.232	0.737	0.633
	(0.259)	(0.232)	(0.208)	(0.287)	(0.259)	(0.233)
R^2	0.115	0.119	0.114	0.115	0.119	0.194
Children Ever Born	2.425	0.781	0.749	2.437	0.873	0.868
	(0.741)	(0.540)	(0.537)	(0.811)	(0.608)	(0.611)
R^2	0.092	0.142	0.147	0.091	0.140	0.146
Employed	-0.820	-0.395	-0.204	-0.934	-0.487	-0.293
	(0.149)	(0.129)	(0.096)	(0.167)	(0.138)	(0.107)
R^2	0.020	0.046	0.204	0.019	0.050	0.204
Weeks Worked	-26.022	-12.338	-3.558	-31.641	-16.851	-7.962
	(7.670)	(7.087)	(5.066)	(8.535)	(7.480)	(5.581)
R^2	0.020	0.041	0.203	0.019	0.043	0.203
Ever Married	0.384	0.377		0.439	0.417	
	(0.119)	(0.122)		(0.127)	(0.127)	
R^2	0.046	0.063		0.046	0.064	
Educ. and Farm Controls	no	yes	yes	no	yes	yes
Marital Status Controls	no	no	yes	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of residence and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state of residence. Regressions for children under age 5 also contain number of children older than 5. Regressions (1) to (3) control for state of birth. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rates are assigned by state of residence. All data are weighted using census person weights (sample line weights for children ever born). For children ever born, the sample is restricted to observation where the number of children does not exceed seven. The sample sizes are 243,554 for columns (1)–(3) and 211,303 for columns (4)–(6) (children ever born: 143,637 for columns (1)–(3) and 121,244 for (4)–(6)).

Table A9: Impact of WWII Mobilization Rates on Labor Supply of Women Aged 45–55 (Coefficient Estimates from OLS Regressions for Variable "Mobilization Rate $\times$ 1960")

Dependent Variable	Regression					
	Full Sample			Stayers Sample		
	(1)	(2)	(3)	(4)	(5)	(6)
Employed	0.058	0.185	0.174	0.099	0.157	0.104
	(0.074)	(0.081)	(0.077)	(0.082)	(0.086)	(0.080)
R^2	0.081	0.112	0.200	0.080	0.112	0.199
Weeks Worked	17.998	17.349	16.552	18.786	15.045	12.171
	(4.384)	(4.326)	(4.641)	(4.651)	(4.403)	(4.492)
R^2	0.066	0.088	0.182	0.064	0.089	0.183
Educ. and Farm Controls	no	yes	yes	no	yes	yes
Marital Status Controls	no	no	yes	no	no	yes

Notes: Standard errors (in parentheses) are adjusted for clusters of state of residence and year of observation. Estimates are from separate regressions of pooled micro data from the 1940 and 1960 censuses. Each outcome variable is regressed on the WWII mobilization rate interacted with a 1960 year indicator variable and indicator variables of observation year, age, race, and state of residence. Regressions (1) to (3) control for state of birth. Where applicable, the farm, race, age, and marital status indicator variables are also interacted with the 1960 indicator variable. Mobilization rates are assigned by state of residence. All data are weighted using census person weights. Sample size is 191,715 for columns (1)–(3) and 179,384 for columns (4)–(6).

birth or the state of current residence as proxies. The census does provide information about whether a woman has moved in the previous five years. Using this information, we construct a sample of "stayers," i.e., women who have not moved in the previous five years. For this sample, the state of residence should be a better approximation for the exposure to the mobilization shock compared to the full sample. Moreover, our main fertility measure considers births over the past five years. Hence, for the stayers sample we can be sure that the current state of residence was the state of residence during this entire fertility period.

Tables A8 and A9 display regression results analogous to columns (4) to (6) of Tables 2 and 4 for the stayers sample. For the young group in Table A8, all parameter estimates are larger for the stayers sample compared to the whole sample, suggesting that focusing on stayers indeed leads to an improved assignment of the mobilization rate and strengthening our overall findings. This is in particular true for the number of children under age 5, where focusing on stayers guarantees that women lived in their assigned state for the entire period when these children were born. For the older group in Table A9, in contrast, the patterns are less clear cut, with some estimates increasing in size while others become smaller in the stayers sample. For the old group, ideally we would like to assign women the mobilization rate of the state where they lived at the end of World War II, and focusing on those who haven't moved between 1955 and 1960 is not a major improvement in this regard.

C Relative Wages for Other Age Groups

In our quantitative analysis, we proxy the relative female wage per efficiency unit by the average wage of single women aged 20–24 relative to the average wage of single men in the same age group. Given that in the data, fertility also rises for other groups, one may wonder how relative wages evolve for older women. Figure A4 displays the relative wages for workers at ages 20–24, 25–29, and 30–34. In each case, the data are for singles, because few married women worked, leading to obvious selection issues for married individuals. As can be seen, the overall pattern is similar for the three groups, with a decline from 1940 to 1950 and 1960 (the baby-boom period) and a subsequent recovery. This is consistent with the interpretation that these women were close substitutes in the labor market. While the patterns thus support the theory, we regard the results for the youngest group as the most reliable, again because of selection effects. After age 24 a large fraction of individuals are married, and marriage rates overall increased substantially over the baby-boom period. It is therefore possible that some of the larger changes for the older women are due to changes in selection. This is less of an issue for the youngest group, because here most individuals are still single. We therefore focus on the 20–24 group as our measure of the gender gap, while noting that the patterns of the older groups are equally in line with the crowding-out interpretation.

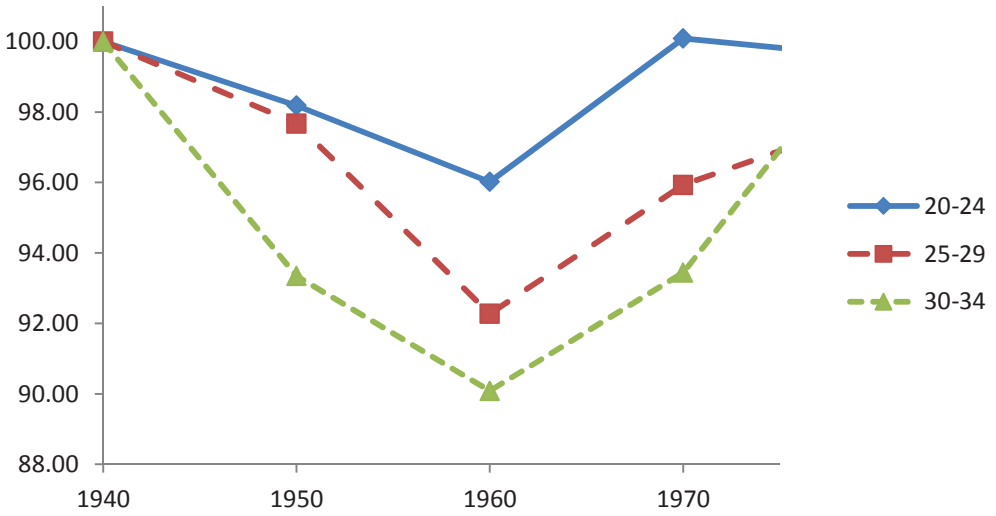

Figure A4: Ratio of Average Female to Average Male Wages for Singles of Different Ages, 1940=100